

Influenza Pandemic
and an Actuarial Model That Takes Account Additional Mortality

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Abstract

Influenza pandemic would be defined as a situation when a new influenza virus emerges and infects humans, starts to spread easily and sustainably among humans, and finally causes widespread epidemic among the world. This is a risk (pandemic risk) that might cause worldwide increase of death and social/economic confusion. This risk has been strongly recognized among medical authorities; however, in the last few years, general public also have heightened their sense of crisis. This article is an effort to address this risk from actuarial perspectives.

In this article, the author first summarizes the general information regarding current understanding of influenza pandemic in the world. Then the author explains the risk characteristics from actuarial point of view by comparing it with some other types of risk (for example, earthquake risk as a catastrophic risk). In later chapter, the author introduces two types of life contingency models which extend traditional actuarial mathematics and which can be used to analyze pandemic and other catastrophe type mortality risks. For whole life policies, the author shows some closed formula of some actuarial notations, such as single premium and net level premium rate. After that, the author considers a life insurance portfolio that consists of n policies. Its present value of loss is examined as a measure of mortality risk of the portfolio. This article shows an approach of calculating a theoretical value of the standard deviation of the loss and it is compared with a stochastic simulation result.

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(Chapter 1) Introduction

(1-1) Influenza pandemic would be defined as a situation when a new influenza virus emerges and infects humans, starts to spread easily and sustainably among humans, and finally causes widespread epidemic among the world. This is a risk (pandemic risk) that might cause worldwide increase of death and social/economic confusion. This risk has been strongly recognized among medical authorities; however, in the last few years, general public also have heightened their sense of crisis. This article is an effort to address this risk from actuarial perspectives.

(1-2) In Chapter 2, the author first summarizes the general information regarding current understanding of influenza pandemic in the world. There are already many well developed reports from governmental organizations and articles written for general public available. By reading these materials, we understand that we face a critical situation regarding the influenza pandemic. In Chapter 3, the author explains risk characteristics from actuarial point of view by comparing it with some other types of risk (for example, earthquake risk as a catastrophic risk). Our main concern in this article is mortality risk, while other risks related to the influenza pandemic is not addressed in this paper.

(1-3) Chapter 4 starts with a brief review of traditional, standard life contingencies, and then we discuss two types of models that incorporate excess mortality caused by pandemic or other catastrophic type mortality risks. These models are extensions of the traditional life contingencies and would be used to analyze these risks. Using whole life policies, some closed formula of some actuarial notations, such as single premium and net level premium rate are obtained. Later in this chapter, we consider an insurance portfolio that consists of n policies. It is well known that, if we assume such excess mortality, the mortality risk of the portfolio is not diversifiable. We examine the present value of loss stochastic variable of the portfolio closely. Taking its standard deviation as a risk measure of the portfolio, this article shows an approach of calculating the theoretical value of the standard deviation in spreadsheets. In Chapter 5, the author developed a stochastic simulation model of the insurance portfolio, and compares the simulated value with the theoretical value determined based on Chapter 4 result. By performing stochastic simulation, we can obtain the distribution of the loss, while the advantage of our theoretical approach is its convenience of calculating the standard deviation. Such easy to use approach can be useful in risk management situation, for example, for a solvency management issues. Finally, Chapter 6 provides some conclusions and remarks.

(1-4) The opinion presented here is solely based on the author's thoughts and would not represent the opinion by the company or the actuarial organization the author belongs to. Also all the errors and misstatement is the responsibility of the author and not of the references.

(Chapter 2) Influenza Pandemic

(2-1) In this chapter, we will look through a basic knowledge of influenza pandemic which might be useful as a background for later chapter discussion. Since there are many well developed reports and articles available, please refer to these sources if you have interested in this subject.

(2-2) What is influenza pandemic?

In this paper, we define influenza pandemic as a situation when a new influenza virus emerges and infects humans, starts to spread easily and sustainably among humans, and finally causes widespread epidemic among the world. This is a risk (pandemic risk) that might cause worldwide increase of death and social/economic confusion.

(2-3) What makes the influenza pandemic as a big threat to our society?

It is predicted that such an emergence and spread of new types of influenza will occur periodically in every several decades. Once it happens, such epidemic might become a world wide phenomenon. Past experience in 1918 of the catastrophic spread of Spanish influenza caused disastrous loss of human lives and we might have a similar or even larger loss in the next pandemic. Such worst case scenario is not a grandiose figment: some type of new influenza virus, called H5N1 type virus, is highly pathogenic. Current development of transportation network has enabled the influenza spread much faster and among the continents. These changes of circumstance tell that the past experience may not be applied in the current situation. We must take in mind that the current situation does not allow optimism.

(2-4) Influenza occurred in the 20th century

There are three influenza pandemics occurred and recorded in the 20th century:

- 1) Spanish Influenza (1918-1919): Total number of the patient counted 25-30% of the total population of the world. Case fatality rate (death over the infected) is more than 2.5%. In Japan, the number of patient is estimated as 23 million and the 380,000 people died ([15]). Virus type is H1N1.
- 2) Asian Influenza (1957-1958): Case fatality rate seems to be much lower than that of Spanish Influenza. Estimated excess number of death is more than 2 million people worldwide. Virus type is H2N2.
- 3) Hong Kong Influenza (1968-1969): The impact is even milder than that of Asian Influenza. Virus type is H3N2.

(2-5) How would the influenza spread among the people and cause pandemic?

There are several stages before new influenza causes pandemic situation. WHO classifies these stages into 6 phases. Following is a simplified explanation of such stages.

- A situation is that there exist new (potential) types of avian influenza in birds (wild birds or poultry); however, there is no infection to human (WHO stage 1 and 2).
- A situation is that infection from birds to human is sometimes reported; however, the infection from human to human is not recognized (WHO stage 3, which is a current situation).
- A situation is that infection from human to human is recognized; however, the infection is still limited and under control (WHO stage 4 and 5)
- A situation is that infection from human to human is done in an efficient and sustained manner and that influenza starts to spread (WHO stage 6, or pandemic)

When an influenza virus obtains a new infection power, genetic mutation occurs in a virus. Such mutations can be either as an antigenetic drift or an antigenetic shift. The former refers to a continuous, evolutionary

mutation process in a virus, while the latter refers to a sudden, evolutionary mutation process. What is known as genetic assortment is a type of antigenetic shift.

(2-6) What is influenza virus? What is highly pathogenic and weakly pathogenic virus?

Influenza virus can be classified into three types, type A, B and C, by the difference of the protein in the virus. Among them, type influenza is known to cause influenza pandemic. Following explanation refers to type A influenza.

On a surface of influenza virus, there exist two types of proteins that outshoot from the surface. One is called hemagglutinin (HA) and the other is called neuraminidase (NA). There are 16 subtypes of HA (H1 to H16) and 9 subtypes of NA (N1 to N9). They have critical roles in infecting and duplication process. The pathogenicity of influenza depends on the HA protein. H5 and H7 subtype is known as highly pathogenic, since once virus with these subtypes infect human, it can attack human organ directly in the body and therefore, infected person tends to show severe symptom and the fatality rate is high. Virus with other HA subtypes are known as low pathogenic. All 20th century pandemics were caused by low pathogenic virus, while the new influenza viruses found in Asian countries recently are H5N1 subtype, which foster the sense of crisis.

(2-7) Possible impact of pandemic

Possible impact of influenza pandemic should be understood from various perspectives. It has impact on human lives, our economy and the society. Also, there are both long-term as well as short-term impacts.

The report by the Infectious Disease Surveillance Center of Japan ([15]) provides the following list of impact. 1) A huge number of patient and death, 2) mental and physical pain, 3) infection to the medical practitioners, 4) excess demand to medical institutions and breakdown of medical services, 5) Infection to the work staffs of social infrastructure (transportation, communication, police, electric, food, water, fire, etc.), 6) a breakdown of social function and public administration, 7) restriction on daily activities, 8) restriction on corporate activities, 9) collapse of the foundation of corporation, due to the domino effect, 10) decrease of population in production age, and 11) a huge economical loss.

The level of impact is affected by various factors. According to the above report, such factor include, 1) pathogenicity of the virus itself, 2) number of infected in epidemic period, 3) proportion of severe complication, 4) Distribution of patients' age, 5) speed of the pandemic spread, 6) the degree of preparation, 7) effect of the action, 8) social background and infrastructure (age distribution of the population, population density, sanitary condition in general, and 9) level of the medical treatment.

With respect to the estimate of human and economic impact, following are some citations from reports prepared by governmental organizations.

<Impact on Human lives>

Ministry of Health, Labour and Welfare of Japan has estimated a number of deaths of the Japanese ([7]). The model they used was originally developed by the Center of Disease Control (CDC) of the U.S.A ([10]). Using that model, it is estimated that the 25% of the total population of Japanese will be infected by the influenza virus. The fatality rate is assumed to be 0.53% for moderate cases (similar to Asian Influenza) and 2% for severe cases (similar to Spanish Influenza). The number of deaths is estimated 170 thousand for moderate cases and 640 thousands for severe cases. The report remarks that effects of the following factors such as intervention with new influenza vaccines or Antiviral drugs and sanitation conditions in Japan are

not taken consideration in these estimates.

<Impact on the Economy >

The reports prepared by the Congressional Budget Office of the U.S.A ([8],[9]) has concluded that the overall impact of a potential flu pandemic on gross domestic product to be about 4.25% in a severe scenario and about 1% in the mild scenario. The reports explain that severe scenario refers to the one similar to Spanish influenza and the mild scenario refers to the one similar to Asian or Hong Kong influenza.

(Chapter 3) Pandemic Risk

(3-1) In this chapter, we discuss the risks that insurers bear when an influenza pandemic occurs. We focus especially on mortality risks.

(3-2) From insurers' perspective, pandemic risk is defined as the risk that occurs when influenza pandemic emerges and the excess claim may cause insurers' insolvency.

(3-3) Pandemic risk may be classified as a kind of catastrophe risk (or calamity risk) like earthquake risk. There are several differences between earthquake and pandemic risk. Following chart summarizes the difference between pandemic and some other risk.

<Chart 1: Comparison of mortality risk>

| Item | Pandemic | Earthquake | War/Terrorism | Life underwriting risk in general |
|---------------------------------|---|--|---------------------------------|---|
| Area | Might be global phenomenon | Local | Assumed Local (*1) | — |
| Time horizon | Several weeks to several years | In a moment to several days | Several days to several decades | — |
| Cycle Length | Once in several decades | Once in several decades per one area | Unpredictable | Death occurs continuously |
| Impact on Human/Property | Human | Both Human/Property | Both Human/Property | Human |
| Impact by age | Might take various patterns | Uniform among ages | Might take various patterns | As shown in mortality table |
| Distribution of loss | Distribution with a long tail | Distribution with a long tail | Unknown | Normal distribution (Law of large numbers) |
| Quantitative evaluation of loss | Difficult (*2) | Some experience in PC insurers | Difficult | Sufficient experience in life insurers |
| Risk diversifiability | Not diversifiable | Not diversifiable | Not diversifiable | Diversifiable |
| Approach to risk | Contingency reserve, Surplus, Reinsurance, Securitization | Contingency (or catastrophe) reserve, Surplus, Reinsurance, Securitization | Exclusion clause | Law of large numbers, Risk selection, Contingency reserve, Surplus, Reinsurance, Securitization |

(*1) Global war is not likely to occur

(*2) Research has just started

<Chart 2> Comparison between pandemic risk and earthquake risk

| | Similar issues | Different issues |
|-----------------|--|--|
| Pandemic risk | <ul style="list-style-type: none"> - Catastrophe (Calamity) risk -Occurs periodically in every several decades - Have a distribution of long-tail - Risk is not diversifiable - Exclusion clause might be applied -Need government level preparation to prevent/ease losses -(Currently in Japan) Contingency reserve I might be used | <ul style="list-style-type: none"> - Might be global (worldwide) in a long term (might be local in a short-term) - The affect may continue for more than several years - Different impact on humans by their ages - Mainly impact on human lives - Early stage counter-action might be critical to prevent /ease the pandemic - Quantitative risk evaluation might be difficult (Model risk and parameter risk might be significant) |
| Earthquake risk | | <ul style="list-style-type: none"> - Limited to local or in a area - Occurs in a very short period (several days) - Impact on all humans in a same degree - Impact on both human lives and property - Some experience of quantitative risk evaluation |

(3-4) As summarized in Chart 2, there are several issues that are not similar between pandemic risk and earthquake risk. First, earthquake risk is limited in relatively local region and the damage is limited in an area of, say, several hundred kilo-meters radius, while for pandemic risk is truly a global phenomenon. In other words, the earthquake risk is geometrically diversifiable while the pandemic risk is not. Once the influenza pandemic spreads all around the world, the impact might be huge.

There are several patterns of how the pandemic affects human by age. If we classify the human into three categories by age, childhood (age 0 to 18), adulthood (19-64), and elderly (65 and above), the impact of pandemic might be different. Following are some typical patterns:

Pattern 1) Childhood and elderly people tends to be affected relatively severely by the influenza while adulthood tends to be affected relatively mildly. This pattern implies that the former classes are thought to be less resistant to influenza than the latter class.

Pattern 2) Adulthood people tends to be affected relatively severely by the influenza while childhood and elderly tends to be affected relatively mildly. One explanation is that once an adult person, who has more resistance to childhood and elderly, comes down with flu, the resistance system of the body overreacts and produce internal substance called cytokine more. Such over-reaction has the opposite effects of making the illness more severe. (This phenomenon is called cytokine storm). Also, since adulthood people contacts with other person much frequently in a society, than childhood and elderly people, they has more opportunity to be affected by the influenza.

Pattern 3) Impact of influenza is same among childhood, adulthood and elderly people. This is the same situation as the impact of earthquake risk.

It is difficult to predict which pattern will emerge. Also, the pattern may be affected by the governmental/social reaction to the pandemic at the beginning of the influenza spread. Temporary closing of school classes and business offices, for example may be effective to prevent the widespread of influenza.

Risk quantification of pandemic risk is difficult and the author believes it is in an early stage, compared to the risk quantification of earthquake risk. For pandemic risk, the actual data is very limited. When we

develop a risk model, model risk and parameter risk might be significantly large and always have to be taken care.

(3-5) Approaches to the mortality risk management

For standard mortality risk, with which we expect that the death occurs randomly and independently, insurance company can diversify the mortality risk by constructing the insurance portfolio as uniform as possible. For instance, insurers may limit the minimum and maximum amount of death benefit, or may apply strict risk selection procedures. This is a traditional approach which utilizes the law of large numbers.

In addition, for a sudden, unexpected risk other than usual death, contingency reserves and surplus funding are in practice to ensure the solvency on both mandatory and voluntary base. Taking the contingency reserve regulation in Japan, life insurers must fund contingency reserve I to cover the mortality risk for ordinary life coverage. Regulatory maximum contingency reserve I requirement for ordinary death is 0.6 permillage of insurance amount at risk. This coefficient might be re-estimated once we have established a well developed method to quantify pandemic risk.

Reinsurance, especially non-proportional type reinsurance such as stop-loss reinsurance can be used to prepare unexpected risk. Regarding the pandemic risk, reinsurance may not work; however, if the pandemic become a global phenomena and the geometrical diversification may not work. As an advanced version of reinsurance, securitization of mortality risk is currently gaining attention.

(3-6) Approaches to risk quantification

To prepare for the pandemic risk, we need any kind of risk quantification. There are several practical, but strong approaches. One approach is a scenario setting, another approach is stochastic simulation. In this paper, the author takes theoretical approach.

Approaches to be taken depend on the risk measure taken for risk quantification. In this paper, we take loss stochastic variable of an insurance portfolio and examine its characteristics. Variance, value at risk (VaR) or tail value at risk (TVaR) are some well known measures for risk quantification. Simulation is the only approach to calculate VaR or TVaR; however, developing sophisticated risk model is time consuming and need lot of resources. This paper's approach is focus mainly on variance, and therefore, there is a limitation; however, it is simple and does not need model development. In practice, several approaches are mixed and used together.

(Chapter 4) Extended Actuarial Model That Reflects Additional Mortality Caused by Pandemic

(4-1) In this section, we examine two actuarial models that include additional mortality caused by a pandemic or some catastrophic events. We start by reviewing some basic results of traditional actuarial mathematics. We see some actuarial expressions for continuous pay whole life. After that, we will extend these results by including additional mortality caused by sudden death by pandemic or other catastrophic events.

<A review of traditional life contingencies>

(4-2) We define some notations as follows;

$T = T(x)$: *stochastic variable that means future lifetime of a life of age (x)*

$F_T(t) = \Pr[T \leq t]$: *Cummulative probability distribution function of T*

$f_T(t) = \frac{\partial F_T(t)}{\partial t}$: *Pr obability density function of T*

$\mu(x+t) = -\frac{1}{F_T(t)} \frac{\partial F_T(t)}{\partial t} = \frac{f_T(t)}{F_T(t)}$: *Force of mortality*

We also define survival function as ${}_t p_x = 1 - F_T(t)$. Using this, we have following expression,

$${}_t p_x = 1 - F_T(t) = \exp\left(-\int_0^t \mu(x+s) ds\right)$$

$$f_T(t) = {}_t p_x \mu(x+t)$$

We assume that a set of continuous mortality is provided a priori by one of the following way; a force of mortality, a cumulative distribution function of T, a probability density function of T, survival function, or other equivalent set of information.

(4-3) We discuss in the continuous time framework; however, if we are given a mortality table $\{q_x\}$ ($x=0,1,2,\dots$), which is a discrete mortality information, we will convert mortality rate of each age into force of mortality by assuming a uniform distribution of death between age $x+t$ and $x+t+1$. In other words, for x and t as integer and for $0 \leq s < 1$, we assume that,

$$\mu(x+t+s) = \frac{q_{x+t}}{1 - sq_{x+t}},$$

And we have,

$${}_s p_{x+t} = 1 - sq_{x+t}$$

(4-4) Take continuous premium pay, immediate benefit pay whole life policy into consideration. Traditional life mathematics is developed by defining some stochastic variables such as Z (present value of future benefit payment) or L (present value of future loss by the company) from T.

Let,

$Z = e^{-\delta T}$: *Stochastic variable that represents present value of future benefit payment*

Then, we can derive some actuarial functions using Z .

- Single premium, present value of life annuities, and variance of Z

$$\bar{A}_x = E[Z] = \int_0^{\omega-x} e^{-\delta t} {}_t p_x \mu(x+t) dt = \frac{i}{\delta} \sum_{t=0}^{w-x} e^{-\delta(t+1)} {}_t q_x = \frac{i}{\delta} A_x$$

$$Var[Z] = {}^{2\delta}\bar{A}_x - (\bar{A}_x)^2$$

$$\bar{a}_x = E[\bar{a}_{\overline{T}|}] = \int_0^{\omega-x} e^{-\delta t} {}_t p_x dt = \frac{1 - \bar{A}_x}{\delta}$$

$$Var[\bar{a}_{\overline{T}|}] = \frac{1}{\delta^2} ({}^{2\delta}\bar{A}_x - (\bar{A}_x)^2)$$

Symbols shown on the upper left of a notation represents a force on interest used for discounting, for example,

$${}^{2\delta}\bar{A}_x = E[\exp(-2\delta T)] = \int_0^{\omega-x} e^{-2\delta t} {}_t p_x \mu(x+t) dt = \frac{e^{2\delta} - 1}{2\delta} \sum_{t=0}^{w-x} e^{-2\delta(t+1)} {}_t q_x = \frac{e^{2\delta} - 1}{2\delta} {}^{2\delta}A_x.$$

2δ means to use 2δ instead of δ to discount future benefit payment. In other words, use discounting factor of $e^{-2\delta t}$ instead of $e^{-\delta t}$ in the integral above.

- Net premium rate

$$\bar{P}_x = \frac{\bar{A}_x}{\bar{a}_x} = \frac{\delta \bar{A}_x}{1 - \bar{A}_x}$$

represents the net premium rate for this whole life policy.

- Future loss stochastic variable (L) and its variance

Using \bar{P}_x , a stochastic variable L is defined by,

$$L = e^{-\delta T} - \bar{P}_x \times \bar{a}_{\overline{T}|}.$$

L represents present value of future loss of insurance company at policy issue. L is the difference between present value of future benefit payment and present value of future premium income. Since net premium is defined using the equivalence principle of profit and loss, $E[L]=0$. Also, we have the following expression,

$$L = \frac{(\delta + \bar{P}_x)Z - \bar{P}_x}{\delta}$$

and the variance of L is shown as,

$$Var[L] = \left(1 + \frac{\bar{P}_x}{\delta}\right)^2 Var[Z] = \left(1 + \frac{\bar{P}_x}{\delta}\right)^2 ({}^{2\delta}\bar{A}_x - (\bar{A}_x)^2).$$

(4-5) Next, we consider an insurance portfolio that consists of n policies, issued at the same time ($t=0$) for n insured, all age (x). Insurance amount is 1. All insured follow the same mortality assumption and death occurs randomly and independently. Termination other than death is not assumed.

It is well known that the mortality risk for this insurance portfolio is diversifiable. Mathematically, this fact is formalized in the following way,

Let present value of benefit payment and present value of future loss for the kth ($k=0,1,2,\dots,n$) policy be Z_k and L_k , respectively. Let the total of L_k be ${}^{Port}L$. ${}^{Port}L$ represents the future loss stochastic variable for the insurance portfolio. We assume that the death occurs independently. In other words, Z_k and Z_l ($k \neq l$) are independent and the covariance of Z_k and Z_l is zero.

Then,

$$L_k = \frac{(\delta + \bar{P}_x)Z_k - \bar{P}_x}{\delta}$$

$$Var[L_k] = \left(1 + \frac{\bar{P}_x}{\delta}\right)^2 Var[Z_k]$$

$$Cov[L_k, L_l] = \left(1 + \frac{\bar{P}_x}{\delta}\right)^2 Cov[Z_k, Z_l] = 0 \quad (k \neq l)$$

$${}^{Port}L = L_1 + L_2 + \dots + L_n$$

$$Var[{}^{Port}L] = nVar[L] + \sum_{k \neq l} Cov[L_k, L_l]$$

$$= n\left(1 + \frac{\bar{P}_x}{\delta}\right)^2 Var[Z]$$

The future loss of the insurance portfolio per policy (${}^{Port}L/n$) converges to zero, as n goes to infinity. In other words, the mortality risk is diversifiable.

$$Var\left[\frac{{}^{Port}L}{n}\right] = \frac{1}{n}\left(1 + \frac{\bar{P}_x}{\delta}\right)^2 Var[Z] \longrightarrow 0 \quad (n \longrightarrow \infty)$$

<Description of the models>

(4-6) Now we consider catastrophe type risks such as pandemic risk or earthquake risk and try to include these risks in traditional life contingencies. In the following two models, it is assumed that additional death caused by a catastrophe is assumed, in addition to usual death assumption which is reflected in mortality table or force of mortality. More specifically, it is assumed that when a catastrophe event occurs, $1-R$ of the total population would die immediately while the rest (R) survives after the event. R might be either a constant value or a stochastic value, depending on the model. Some detailed assumptions are added to have the following two types of the models (Model A and Model B). In developing these models, the author refers to Hu and Cox [4] for their ideas.

(4-7) Model A

In model A, we assume that a catastrophic event occurs once in $1/\lambda$ years in average. When the event occurs, $1-r$ of the population dies and the rest (r) survives. Here λ and r are assumed to be a constant. Catastrophic events occur not only once. The event occurs with the same probability in the future. In other words, the interval stochastic variable, U_1, U_2, \dots between two the events follows an exponential distribution with average $1/\lambda$. Mathematically written, cumulative probability distribution functions are shown as,

$$F_U(u) = 1 - \exp(-\lambda u)$$

where U_1, U_2, \dots are independent each other.

It is well known that, if there are K catastrophic event occurred in time 0 to t , K follows Poisson distribution with average λt ;

$$\Pr[K = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}.$$

(4-8) Model B

In model B, catastrophic event occurs only once. It assumes to occur in $1/\lambda$ years in average. When the event occurs, $1-R$ of the population dies and the rest (R) survives. Here λ is a constant but R is assumed to be a stochastic variable. R takes a value between 0 and 1, and $E[R]=r$, $\text{Var}[R]=\sigma^2$. Stochastic variable U that represents the time of stochastic event follows exponential distribution with average $1/\lambda$.

<Closed expressions of some actuarial functions>

(4-9) For model A, some actuarial expressions such as single premium are obtained in a closed form. In the following notations, star signs (*A,*B) on the upper right implies model A or model B (A,B may be omitted).

Let $T^{*A}(x)$ be future lifetime stochastic variable in model A. We will obtain cumulative probability distribution function, $F_{T^*}(t)$. Assume that catastrophic event occurs K ($=0,1,2,\dots$) times until time t ,

$$F_{T^*}(t) = \Pr[T^* \leq t] = 1 - \sum_{k=0}^{\infty} \Pr[T^* > t | K = k] \times \Pr[K = k].$$

Here $\Pr[T^* > t | K = k]$ implies the probability that the insured survives at time t , assuming that the event occurred k times,

$$\Pr[T^* > t] = r^k \times {}_t p_x.$$

Since K follows Poisson distribution with average λt ,

$$\begin{aligned} F_{T^*}(t) &= 1 - \sum_{k=0}^{\infty} r^k \times {}_t p_x \times e^{-\lambda t} \frac{(\lambda t)^k}{k!} \\ &= 1 - {}_t p_x \times e^{-\lambda t} \sum_{k=0}^{\infty} \frac{(r\lambda t)^k}{k!} \\ &= 1 - {}_t p_x \times e^{-\lambda(1-r)t} \end{aligned}$$

Therefore,

$$\mu^*(x+t) = \mu(x+t) + \lambda(1-r)$$

$$f_{T^*}(t) = {}_t p_x \times \mu(x+t) \times e^{-\lambda(1-r)t} + \lambda(1-r) \times {}_t p_x \times e^{-\lambda(1-r)t}$$

It is worth to comment on the expression of force of mortality shown above. This expression implies that the force of mortality is given by an assumed mortality $\mu(x+t)$ plus additional constant mortality that represents the probability of catastrophic event. Compare with a situation of additional constant mortality

for an insured that has hazard occupation such as a pilot. He or she faces additional mortality, which threatens his or her life continuously. The expression of force of mortality doesn't distinguish the nature of the risk. In other words, we can't judge the nature of the risk by just looking at the expression of the force of mortality. We will see later that the additional mortality for hazard occupation is diversifiable while that for catastrophe is not diversifiable.

(4-10) Next, define Z^* as, $Z^* = \exp(-\delta T^*)$, and

$$\begin{aligned}\bar{A}_x^* &= E[Z^*] \\ &= \int_0^{\omega-x} e^{-\delta t} \times f_{T^*}(t) dt \\ &= \int_0^{\omega-x} e^{-\delta t} \times ({}_t p_x \times \mu(x+t) \times e^{-\lambda(1-r)t} + \lambda(1-r) \times {}_t p_x \times e^{-\lambda(1-r)t}) dt \\ &= \int_0^{\omega-x} e^{-(\delta+\lambda(1-r))t} \times {}_t p_x \times \mu(x+t) dt + \lambda(1-r) \int_0^{\omega-x} e^{-(\delta+\lambda(1-r))t} \times {}_t p_x dt\end{aligned}$$

Here we introduce the following notation,

$$\begin{aligned}{}^{\delta+\lambda(1-r)}\bar{A}_x &= \int_0^{\omega-x} e^{-(\delta+\lambda(1-r))t} \times {}_t p_x \times \mu(x+t) dt \\ {}^{\delta+\lambda(1-r)}\bar{a}_x &= \int_0^{\omega-x} e^{-(\delta+\lambda(1-r))t} \times {}_t p_x dt\end{aligned}$$

This notation represents the single premium for a whole life for $T(x)$, replacing the force of interest from

δ to $\delta + \lambda(1-r)$. Note that ${}^\delta\bar{A}_x = \bar{A}_x$, ${}^\delta\bar{a}_x = \bar{a}_x$.

Since we have the following relation,

$$1 = ({}^{\delta+\lambda(1-r)}\bar{A}_x) + (\delta + \lambda(1-r)) ({}^{\delta+\lambda(1-r)}\bar{a}_x)$$

we have,

$$\begin{aligned}\bar{A}_x^* &= {}^{\delta+\lambda(1-r)}\bar{A}_x + \lambda(1-r) \times {}^{\delta+\lambda(1-r)}\bar{a}_x \\ &= 1 - \delta \times {}^{\delta+\lambda(1-r)}\bar{a}_x \\ &= \frac{1}{\delta + \lambda(1-r)} (\lambda(1-r) + \delta \times {}^{\delta+\lambda(1-r)}\bar{A}_x)\end{aligned}$$

Also, $1 = \bar{A}_x^* + \delta \times \bar{a}_x^*$ and

$$\bar{a}_x^* = {}^{\delta+\lambda(1-r)}\bar{a}_x.$$

We have the expression of the variance of the present value of benefit payment as,

$$\begin{aligned}
Var[Z^*] &= E[(Z^*)^2] - (E[Z^*])^2 \\
&= 2\delta + \lambda(1-r)\bar{A}_x + \lambda(1-r)\delta \times 2\delta + \lambda(1-r)\bar{a}_x - (\bar{A}_x)^2 \\
&= \dots \\
&= (1 - 2\delta \times 2\delta + \lambda(1-r)\bar{a}_x) - (1 - \delta \times \delta + \lambda(1-r)\bar{a}_x)^2 \\
&= \frac{1}{2\delta + \lambda(1-r)} \left(\lambda(1-r) + 2\delta \times 2\delta + \lambda(1-r)\bar{A}_x \right) - \left(\frac{1}{\delta + \lambda(1-r)} \left(\lambda(1-r) + \delta \times \delta + \lambda(1-r)\bar{A}_x \right) \right)^2
\end{aligned}$$

Also, if we define the present value of future loss stochastic variable as, $L^* = \exp(-\delta T^*) - \bar{P}_x^* \times \bar{a}_{\overline{T^*}|}$,

where \bar{P}_x^* be a net premium rate,

$$\bar{P}_x^* = \frac{\bar{A}_x^*}{\bar{a}_x^*} = \frac{\delta \bar{A}_x^*}{1 - \bar{A}_x^*}$$

we have the variance of loss stochastic variable as,

$$Var[L^*] = \left(1 + \frac{\bar{P}_x^*}{\delta}\right)^2 \times Var[Z^*].$$

(4-10) Also for model B, we can express some actuarial expressions such as single premium in a closed form. First, we consider the situation where $R=r$ (constant). In a similar way as model A,

$$F_{T^*}(t) = \Pr[T^* \leq t] = \sum_{k=0}^1 \Pr[T^* \leq t | K = k] \times \Pr[K = k].$$

Here $K=0$ implies that the event has not been occurred yet, while $K=1$ implies that the event has already occurred.

$$\begin{aligned}
F_{T^*}(t) &= \Pr[T^* \leq t | K = 0] \times \Pr[K = 0] + \Pr[T^* \leq t | K = 1] \times \Pr[K = 1] \\
&= (1 - {}_t p_x) \times e^{-\lambda t} + (1 - r \times {}_t p_x) \times (1 - e^{-\lambda t}) \\
&= 1 - {}_t p_x \times (r + (1-r)e^{-\lambda t})
\end{aligned}$$

For a random variable R in general, let the probability density function of R as, $g_R(s)$ and let $E[R] = r$,

$$\begin{aligned}
F_{T^*}(t) &= \int_0^1 g_R(s) \times (1 - {}_t p_x \times (s + (1-s)e^{-\lambda t})) ds \\
&= 1 - {}_t p_x \times (r + (1-r)e^{-\lambda t})
\end{aligned}$$

The cumulative probability distribution function has the same form. In other words, above expression does not rely on the distribution of R , but only rely on the expected value of R ($=r$). By taking the derivative in the right side, we have,

$$f_{T^*}(t) = {}_t p_x \times \mu_{x+t} \times (r + (1-r)e^{-\lambda t}) + {}_t p_x \times (1-r)\lambda e^{-\lambda t}$$

Then we have,

$$\begin{aligned}\bar{A}_x^* &= r\bar{A}_x + (1-r)^{(\delta+\lambda)}\bar{A}_x + \lambda(1-r)\times^{(\delta+\lambda)}\bar{a}_x \\ &= r\bar{A}_x + (1-r)\times\left(\frac{\lambda}{\delta+\lambda} + \frac{\delta}{\delta+\lambda}^{(\delta+\lambda)}\bar{A}_x\right)\end{aligned}$$

Using the relation, $1 = \bar{A}_x^* + \delta \times \bar{a}_x^*$,

$$\bar{a}_x^* = r \times \bar{a}_x + (1-r)\times^{(\delta+\lambda)}\bar{a}_x$$

For reference, above single premium formula can be obtained in the following way. First, assume that a catastrophic event occurs at time τ , which is a constant value. “ r ” is assumed to be constant for a time being.

$$\begin{aligned}\bar{A}_x^* &= \bar{A}_{x:\tau}^1 + e^{-\delta\tau} \times {}_r p_x ((1-r) + r\bar{A}_{x+\tau}) \\ &= (1-r)\bar{A}_{x:\tau} + r\bar{A}_x\end{aligned}$$

Above expression is a weighted average of a single premium of endowment life and a single premium of a whole life. When τ is not a constant but has the exponential distribution with average $1/\lambda$, the first term turns out to be,

$$\begin{aligned}\int_0^\infty \bar{A}_{x:\tau} \times \lambda e^{-\lambda\tau} d\tau &= \Lambda \\ &= \frac{\lambda}{\delta+\lambda} + \frac{\delta}{\delta+\lambda}^{(\delta+\lambda)}\bar{A}_x\end{aligned}$$

and we have the expression in (5-10). Same argument is valid when r is a random variable R .

(4-11) We have closed expressions for $\text{Var}[Z^*]$ and $\text{Var}[L^*]$.

$$\begin{aligned}\text{Var}[Z^*] &= E[(Z^*)^2] - (E[Z^*])^2 \\ &= r \times {}^{2\delta}\bar{A}_x + (1-r)\times\left(\frac{\lambda}{2\delta+\lambda} + \frac{2\delta}{2\delta+\lambda}^{(2\delta+\lambda)}\bar{A}_x\right) - \left(r \times \bar{A}_x + (1-r)\times\left(\frac{\lambda}{\delta+\lambda} + \frac{\delta}{\delta+\lambda}^{(\delta+\lambda)}\bar{A}_x\right)\right)^2\end{aligned}$$

$$\bar{P}_x^* = \frac{\bar{A}_x^*}{\bar{a}_x^*} = \frac{\delta \bar{A}_x^*}{1 - \bar{A}_x^*}$$

$$\text{Var}[L^*] = \left(1 + \frac{\bar{P}_x^*}{\delta}\right)^2 \times \text{Var}[Z^*]$$

<Numerical example>

(4-13) Using the 2007 Japan standard mortality table for male, whole life single premium, present value of life annuities and whole life net premium rate are calculated and shown in Appendix A. The author assumed that $\lambda=0.02$, $E[R]=r=0.98$, $i=2\%$. In calculation, uniform distribution of death assumption for

fractional years is assumed and the formula, $\bar{A}_x = \frac{i}{\delta} A_x$, is used. For age, $X=30, 50, 70$, the result is

shown in the following table.

<Table 1> Comparison of Net Premium Rate (Formula base)

Mortality Table: 2007 Japan Standard Mortality Table (Male)
Assumed Interest Rate: $i=2\%$

Net Premium Rate

| Age | 30 | 50 | 70 |
|-----------|---------|---------|---------|
| Std Model | 0.01269 | 0.02522 | 0.06383 |
| Model A | 0.01296 | 0.02548 | 0.06411 |
| Model B | 0.01289 | 0.02543 | 0.06408 |

Single Premium

| Age | 30 | 50 | 70 |
|-----------|---------|---------|---------|
| Std Model | 0.39047 | 0.55061 | 0.76323 |
| Model A | 0.39562 | 0.55331 | 0.76400 |
| Model B | 0.39434 | 0.55281 | 0.76392 |

Relative value (Std model =100%)

| Age | 30 | 50 | 70 |
|-----------|---------|---------|---------|
| Std Model | 100.00% | 100.00% | 100.00% |
| Model A | 102.18% | 101.06% | 100.43% |
| Model B | 101.64% | 100.86% | 100.38% |

Relative value (Std model =100%)

| Age | 30 | 50 | 70 |
|-----------|---------|---------|---------|
| Std Model | 100.00% | 100.00% | 100.00% |
| Model A | 101.32% | 100.49% | 100.10% |
| Model B | 100.99% | 100.40% | 100.09% |

For $X=30$ male, the author also performed simulation with $N=21,000$ and simulation trial 10,800 times to find simulated results, to confirm the theoretical figures. As shown below, the theoretical calculation has very close result to the simulated result, as expected.

<Table 2> Comparison of Net Premium Rate (Formula vs. Simulation)

Age 30

| | Standard Model | | | Model A | | | Model B | | |
|------------------|----------------|---------|---------|------------|---------|---------|------------|---------|---------|
| | Formula(1) | Sim(2) | (2)/(1) | Formula(1) | Sim(2) | (2)/(1) | Formula(1) | Sim(2) | (2)/(1) |
| Single Premium | 0.39047 | 0.39046 | 99.998% | 0.39562 | 0.39561 | 99.998% | 0.39434 | 0.39434 | 99.999% |
| Net Premium Rate | 0.01269 | 0.01269 | 99.997% | 0.01296 | 0.01296 | 99.997% | 0.01289 | 0.01289 | 99.998% |

Additional net premium with extra mortality by catastrophic event, expressed in relative to that for standard model does not look significant. Premium based on model A is 102.18% of the premium based on standard model. Such small increase is, in reality, might be covered by medical selection. As we see later; however, that the mortality risk measured by standard deviation is much larger in model A or B, compared to the standard model.

<Non-diversifiable mortality risk of insurance portfolio>

(4-14) Next we examine an insurance portfolio of n contacts for model A and model B. We will see that the mortality risk is not diversifiable in this portfolio. As in (4-5), we assume that n policies of continuous pay whole life are issued at the same time ($t=0$) for all age (x). Insurance amount is 1. All insured follow the same mortality assumption (including catastrophic) and death occurs randomly. Termination other than death is not assumed. We examine the expected value of future loss as follows.

(4-15) First we assume model A. In this case, insurance event (death by the insured) occurs randomly based on mortality table and the occurrence of the catastrophic event. The insurance event; however, is not independent, since catastrophic event affects all the insured similarly and at the same time. There is a correlation between the deaths by the insured.

(4-16) For an insurance portfolio with n insured, let Z_k^* and L_k^* be present value of benefit payment and

present value of future loss for k th insured ($k=1,2,\dots,n$) . Let the sum of L_k^* be ${}^{Port}L^*$. We have the following formula,

$$L_k^* = \frac{(\delta + \bar{P}_x^*)Z_k^* - \bar{P}_x^*}{\delta}$$

$$Var[L_k^*] = (1 + \frac{\bar{P}_x^*}{\delta})^2 Var[Z_k^*]$$

$$Cov[L_k^*, L_l^*] = (1 + \frac{\bar{P}_x^*}{\delta})^2 Cov[Z_k^*, Z_l^*] \quad (k \neq l).$$

$${}^{Port}L^* = L_1^* + L_2^* + \dots + L_n^*$$

$$Var[{}^{Port}L^*] = nVar[L_k^*] + \sum_{k \neq l} Cov[L_k^*, L_l^*]$$

We can calculate $Var[{}^{Port}L^*]$ experimentally by developing a simulation model, as shown later. Also if we can calculate either $Cov[L_k^*, L_l^*]$ or $Cov[Z_k^*, Z_l^*]$ in some way, we can calculate $Var[{}^{Port}L^*]$ in a theoretical way, using spreadsheets. In order to do that, we will determine joint cumulative probability distribution function of two lifetime stochastic variables as follows.

Let $T_1^*, T_2^*, \dots, T_n^*$ be n lifetime stochastic variables of insured aged (x) , a joint probability density function of two of these variables, T_1^*, T_2^* is expressed as,

$$F_{T_1^*, T_2^*}(t_1, t_2) = \Pr[T_1^* \leq t_1, T_2^* \leq t_2]$$

First we assume $t_1 \leq t_2$. If there are U catastrophic event happened in time 0 to time t_1 , and there are V catastrophic event happened in time t_1 and t_2 , U and V follow Poisson distribution with average λt_1 and $\lambda(t_2-t_1)$ respectively, and we have,

$$\begin{aligned} F_{T_1^*, T_2^*}(t_1, t_2) &= \sum_{u=0}^{\infty} \sum_{v=0}^{\infty} (1 - r^{u+v} p_x) (1 - r^{u+v} p_x) \times e^{-\lambda t_1} \frac{(\lambda t_1)^u}{u!} \times e^{-\lambda(t_2-t_1)} \frac{(\lambda(t_2-t_1))^v}{v!} \\ &= 1 - p_x \sum_{u=0}^{\infty} e^{-\lambda t_1} \frac{(\lambda r t_1)^u}{u!} - p_x \sum_{u=0}^{\infty} e^{-\lambda t_1} \frac{(\lambda r t_1)^u}{u!} \sum_{v=0}^{\infty} e^{-\lambda(t_2-t_1)} \frac{(\lambda r(t_2-t_1))^v}{v!} \\ &\quad + p_x p_x \sum_{u=0}^{\infty} e^{-\lambda t_1} \frac{(\lambda r^2 t_1)^u}{u!} \sum_{v=0}^{\infty} e^{-\lambda(t_2-t_1)} \frac{(\lambda r(t_2-t_1))^v}{v!} \\ &= 1 - p_x e^{-\lambda(1-r)t_1} - p_x e^{-\lambda(1-r)t_2} + p_x p_x \exp(-\lambda t_1 + \lambda r^2 t_1 - \lambda(t_2-t_1) + \lambda r(t_2-t_1)) \\ &= (1 - p_x e^{-\lambda(1-r)t_1}) (1 - p_x e^{-\lambda(1-r)t_2}) + p_x p_x e^{-\lambda(1-r)(t_1+t_2)} (e^{\lambda t_1 \lambda(1-r)^2} - 1) \end{aligned}$$

Also consider the situation of $t_1 > t_2$ and consider symmetric property, we have,

$$F_{T_1^*, T_2^*}(t_1, t_2) = (1 - p_x e^{-\lambda(1-r)t_1}) (1 - p_x e^{-\lambda(1-r)t_2}) + p_x p_x e^{-\lambda(1-r)(t_1+t_2)} (e^{\lambda \min(t_1, t_2) \lambda(1-r)^2} - 1).$$

The second term implies the correlation between T_1^* and T_2^* , and we confirmed that that T_1^* and T_2^* are not independent.

In addition, above formula is also valid if we assume that the age of the first insured is (x) and that of the second (y), where (x) and (y) might be different. Therefore, we have

$$F_{T_1^*, T_2^*}(t_1, t_2) = (1 - {}_{t_1}p_x e^{-\lambda(1-r)t_1})(1 - {}_{t_2}p_y e^{-\lambda(1-r)t_2}) + {}_{t_1}p_x {}_{t_2}p_y e^{-\lambda(1-r)(t_1+t_2)} (e^{\lambda \text{Min}(t_1, t_2)(1-r)} - 1).$$

(4-17) Once we have joint cumulative probability distribution function, we might calculate the joint probability density function and can calculate the covariance of Z_1, Z_2 . Although joint cumulative probability distribution function is continuous but not differentiable when $t_1=t_2$, following alternative approach can be taken. Discrete version of the joint probability density function can be defined for $m_1=1, 2, \dots, m_2=1, 2, \dots$,

$$\begin{aligned} f_{T_1^*, T_2^*}(m_1, m_2) &= \Pr(m_1 - 1 \leq T_1^* < m_1, m_2 - 1 \leq T_2^* < m_2) \\ &= F_{T_1^*, T_2^*}(m_1, m_2) + F_{T_1^*, T_2^*}(m_1 - 1, m_2 - 1) - F_{T_1^*, T_2^*}(m_1 - 1, m_2) - F_{T_1^*, T_2^*}(m_1, m_2 - 1) \end{aligned}$$

Using this formula and assuming the uniform deaths occurs at the fraction of years, we have,

$$\text{Cov}[Z^*_1, Z^*_2] = \left(\frac{i}{\delta}\right)^2 \sum_{m_1=1}^{\infty} \sum_{m_2=1}^{\infty} e^{-\delta(m_1+m_2)} f_{T_1^*, T_2^*}(m_1, m_2),$$

which can be used to calculate the covariance using spreadsheets.

(4-18) Using the covariance calculated as above, we can calculate the variance of the insurance portfolio defined in (4-16),

$$\begin{aligned} \text{Var}[\text{Port } L^*] &= \left(1 + \frac{\bar{P}_x^*}{\delta}\right)^2 (n \text{Var}[Z^*] + n(n-1) \text{Cov}[Z^*_1, Z^*_2]) \\ \text{Var}\left[\frac{\text{Port } L^*}{n}\right] &= \left(1 + \frac{\bar{P}_x^*}{\delta}\right)^2 \left(\frac{1}{n} \text{Var}[Z^*] + \left(1 - \frac{1}{n}\right) \text{Cov}[Z^*_1, Z^*_2]\right) \end{aligned}$$

Above expression tells that the variance of portfolio loss per contract does not converge to zero for significantly large n. In other words, Portfolio loss is not diversifiable.

For large n, we have an approximation,

$$\text{Var}\left[\frac{\text{Port } L^*}{n}\right] \approx \left(1 + \frac{\bar{P}_x^*}{\delta}\right)^2 \text{Cov}[Z^*_1, Z^*_2].$$

If divided by premium rate, we have the following expression.

$$\begin{aligned} \frac{1}{\bar{P}_x^*} \text{Std}\left[\frac{\text{Port } L^*}{n}\right] &\approx \left(\frac{1}{\bar{P}_x^*}\right) \sqrt{\text{Cov}[Z^*_1, Z^*_2]} \\ &= \left(\frac{1}{\delta}\right) \sqrt{\frac{\text{Cov}[Z^*_1, Z^*_2]}{E[Z^*_1]E[Z^*_2]}} = \left(\frac{1}{\delta}\right) \sqrt{\text{Corr}[Z^*_1, Z^*_2]} \end{aligned}$$

(4-19) Also for model B, we can calculate joint cumulative probability distribution function in a closed form and we can calculate covariance. Let $E[R]=r$, and $\text{Var}[R]=r$. Let the probability density function of R be $g_R(s)$. Assume that $t_1 \leq t_2$. We have the following formula, by considering when the catastrophe occurs

either before t_1 , sometime between t_1 and t_2 , or after t_2 ,

$$\begin{aligned}
& F_{T_1^*, T_2^*}(t_1, t_2) \\
&= \int_0^1 g_R(s) \left((1-s_{t_1} p_x)(1-s_{t_2} p_x)(1-e^{-\lambda t_1}) + (1-{}_{t_1} p_x)(1-s_{t_2} p_x)(e^{-\lambda t_1} - e^{-\lambda t_2}) + (1-{}_{t_1} p_x)(1-{}_{t_2} p_x)(e^{-\lambda t_2}) \right) \\
&= \int_0^1 g_R(s) \left((1-s_{t_1} p_x)(1-s_{t_2} p_x) - (1-s)({}_{t_1} p_x e^{-\lambda t_1} + {}_{t_2} p_x e^{-\lambda t_2}) + (1-s) {}_{t_1} p_x {}_{t_2} p_x (s e^{-\lambda t_1} + e^{-\lambda t_2}) \right) \\
&= (1-r_{{}_{t_1} p_x})(1-r_{{}_{t_2} p_x}) + \sigma^2 {}_{t_1} p_x {}_{t_2} p_x - (1-r)({}_{t_1} p_x e^{-\lambda t_1} + {}_{t_2} p_x e^{-\lambda t_2}) \\
&\quad + {}_{t_1} p_x {}_{t_2} p_x ((r-r^2-\sigma^2)e^{-\lambda t_1} + (1-r)e^{-\lambda t_2}) \\
&= (1-r_{{}_{t_1} p_x})(1-r_{{}_{t_2} p_x}) - (1-r)({}_{t_1} p_x e^{-\lambda t_1} + {}_{t_2} p_x e^{-\lambda t_2}) + (1-r) {}_{t_1} p_x {}_{t_2} p_x (r e^{-\lambda t_1} + e^{-\lambda t_2}) \\
&\quad + \sigma^2 {}_{t_1} p_x {}_{t_2} p_x (1-e^{-\lambda t_1})
\end{aligned}$$

Also considering the situation that $t_1 < t_2$, and the case if the age of the second insured is (y), we have,

$$\begin{aligned}
& F_{T_1^*, T_2^*}(t_1, t_2) \\
&= (1-r_{{}_{t_1} p_x})(1-r_{{}_{t_2} p_y}) - (1-r)({}_{t_1} p_x e^{-\lambda t_1} + {}_{t_2} p_y e^{-\lambda t_2}) + (1-r) {}_{t_1} p_x {}_{t_2} p_y (r e^{-\lambda \text{Min}(t_1, t_2)} + e^{-\lambda \text{Max}(t_1, t_2)}) \\
&\quad + \sigma^2 {}_{t_1} p_x {}_{t_2} p_y (1-e^{-\lambda \text{Min}(t_1, t_2)})
\end{aligned}$$

Above expression tells that the joint cumulative probability distribution function reflects expected value and variance of R. The distribution of R is not relevant. Also, if $\sigma^2=0$, the fourth term becomes zero.

For model B, insurance risk caused by catastrophic event is not diversifiable, using the same argument as in (5-17) and (5-18).

(Chapter 5) Distribution of Present Value of Future Loss by Stochastic Simulation

(5-1) In the previous chapter, we define two models and get the closed expressions of some actuarial notation. Also, for insurance portfolio with n contracts, we get a theoretical value of the variance of present value of future loss by determining the joint cumulative probability distribution function and using it in the discrete version.

We know that insurance risk for catastrophic event is not diversifiable and we expect that actual distribution of future loss has a distribution with long tail in the left side. In order to obtain such distribution, we need to perform stochastic simulation. Following is an example of such stochastic simulation for model A and model B. Our approach can be used to check the validity of the simulation models. Simulation result is compared to the theoretical values calculated using above approach.

(5-2) Stochastic simulation assumption

- Mortality table: 2007 Japan standard mortality table (For fractional years, uniform deaths assumption is used)
- Assumed interest: $i=2\%$
- An insurance portfolio of $x=30$, male, $n=21,000$, of continuous premium pay, immediate benefit pay whole life, all contracted simultaneously at time 0
- Termination other than death is not assumed.
- Insurance amount, 10 million yen per policy (= crude approximation, \$100,000 per policy)
- Simulation 10,800 times
- Model parameters: as follows

<Table 3> Simulation Parameters

| | | |
|------|----------------|--|
| SIM1 | Standard model | $\lambda=0.00$, $R=r=1$, i.e. standard life contingency model |
| SIM2 | Model A | $\lambda=0.02$, $R=r=0.98$ (constant) |
| SIM3 | Model B1 | $\lambda=0.02$, $R=r=0.98$ (constant) |
| SIM4 | Model B2 | $\lambda=0.02$, R has the following probability density function $g_R(s)=49s^{48}$, $E[R]=0.98$ |

For simulation, the author assumes parameters as $\lambda=0.02$ and $r=0.98$, considering a possible worst case scenario. Also for model B2, power function is used for the distribution of R .

(5-3) Simulation result and observations

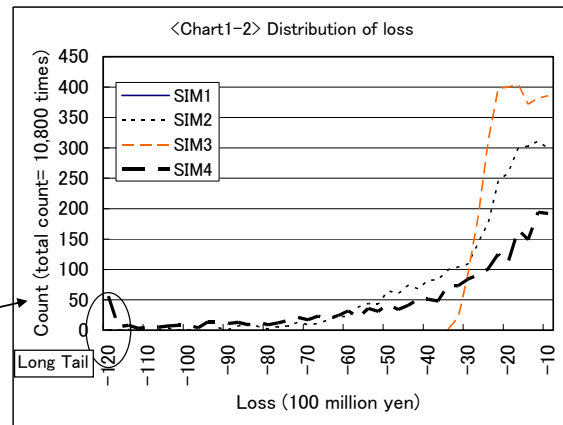
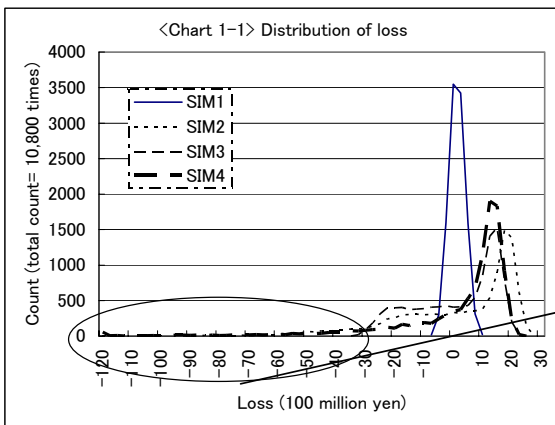
First, a simulation result and the comparison with theoretical calculation are shown as Table 4. For a distribution of each simulation, readers may see Appendix B. As expected, variance of the insurance portfolio is close to each other. The difference is in the 1% range, and often much smaller than 1%.

By stochastic simulation, we see that the portfolio risk has a distribution with long tail, as shown in Chart 1. Generally speaking, in order to quantify the mortality risk inherited in insurance portfolio, we may use several risk measures, such as standard deviation, value at risk (VaR) and tail value at risk (TVaR). As a risk measure, standard deviation has a shortage of not incorporating the asymmetric risk distribution or risk with long tail. But as we have shown, we can calculate standard deviation measure much easier than performing simulation. In Appendix C, standard deviation of insurance portfolio per policy for all x is shown, assuming the n is large enough. Such convenience can be utilized, for example, to perform sensitivity analysis.

<Table 4> Portfolio risk comparison (Formula bs Simulation)

| | Symbol | (SIM1) Standard Model | | | (SIM2) Model A | | |
|---------------------------------------|----------|-----------------------|-------------|---------|----------------|---------------|---------|
| | | Formula(1) | Sim(2) | (2)/(1) | Formula(1) | Sim(2) | (2)/(1) |
| N=21,000, Insured amount = 1 | | | | | | | |
| Expected value of loss | E[L] | 0.00000 | 0.00000 | - | 0.00000 | 0.00000 | - |
| Variance of loss | Var[L] | 719.66392 | 711.03772 | 98.80% | 45,290.30538 | 44,873.24145 | 99.08% |
| Standard deviation of loss | Std[L] | 26.82655 | 26.66529 | 99.40% | 212.81519 | 211.83305 | 99.54% |
| Variance of loss per policy | Var[L/N] | 0.03427 | 0.03386 | 98.80% | 2.15668 | 2.13682 | 99.08% |
| Standard deviation of loss per policy | Std[L/N] | 0.00128 | 0.00127 | 99.40% | 0.01013 | 0.01009 | 99.54% |
| N=21,000, Insured amount = 10 mil yen | | | | | | | |
| Expected value of loss | E[L] | 0 | 0 | - | 0 | 0 | - |
| Variance of loss | Var[L] | 7.19664E+16 | 7.11038E+16 | 98.80% | 4.52903E+18 | 4.48732E+18 | 99.08% |
| Standard deviation of loss | Std[L] | 268,265,525 | 266,652,905 | 99.40% | 2,128,151,907 | 2,118,330,509 | 99.54% |
| Variance of loss per policy | Var[L/N] | 3.42697E+12 | 3.38589E+12 | 98.80% | 2.15668E+14 | 2.13682E+14 | 99.08% |
| Standard deviation of loss per policy | Std[L/N] | 12,775 | 12,698 | 99.40% | 101,341 | 100,873 | 99.54% |

| | Symbol | (SIM2) Model B1 | | | (SIM2) Model B2 | | |
|---------------------------------------|----------|-----------------|---------------|---------|-----------------|---------------|---------|
| | | Formula(1) | Sim(2) | (2)/(1) | Formula(1) | Sim(2) | (2)/(1) |
| N=21,000, Insured amount = 1 | | | | | | | |
| Expected value of loss | E[L] | 0.00000 | 0.00000 | - | 0.00000 | 0.00000 | - |
| Variance of loss | Var[L] | 19,950.06964 | 19,929.35301 | 99.90% | 55,641.62648 | 55,635.59084 | 99.99% |
| Standard deviation of loss | Std[L] | 141.24472 | 141.17136 | 99.95% | 235.88477 | 235.87198 | 99.99% |
| Variance of loss per policy | Var[L/N] | 0.95000 | 0.94902 | 99.90% | 2.64960 | 2.64931 | 99.99% |
| Standard deviation of loss per policy | Std[L/N] | 0.00673 | 0.00672 | 99.95% | 0.01123 | 0.01123 | 99.99% |
| N=21,000, Insured amount = 10 mil yen | | | | | | | |
| Expected value of loss | E[L] | 0 | 0 | - | 0 | 0 | - |
| Variance of loss | Var[L] | 1.99501E+18 | 1.99294E+18 | 99.90% | 5.56416E+18 | 5.56356E+18 | 99.99% |
| Standard deviation of loss | Std[L] | 1,412,447,154 | 1,411,713,605 | 99.95% | 2,358,847,737 | 2,358,719,798 | 99.99% |
| Variance of loss per policy | Var[L/N] | 9.50003E+13 | 9.49017E+13 | 99.90% | 2.6496E+14 | 2.64931E+14 | 99.99% |
| Standard deviation of loss per policy | Std[L/N] | 67,259 | 67,224 | 99.95% | 112,326 | 112,320 | 99.99% |



<Table 5> Standard deviation of loss per policy, divided by net premium rate

| | Symbol | Std Model | Model A1 | Model B1 | Model B2 |
|---------------------------------------|---------------------------------|-----------|----------|----------|----------|
| Standard deviation of loss per policy | $(1/P \cdot x) \text{Std}[L/N]$ | 0.101 | 0.782 | 0.522 | 0.871 |

In order to gasp the level of risk, standard deviation per policy might be standardized by, for example, dividing it with net premium rate. Table 5 shows the standard deviation divided by net premium rate. We see that the standard deviation of loss in the standard model is about 0.1 years of premium, while that in the model B2 is 0.87 years of premium. Under model B2, if the company want to cover the loss with the 99.3% of probability, the company needs to have 2.5 times of standard deviation, or in this case, 2.18 (0.871 * 2.5) years of net premium as surplus.

(Chapter 6) Conclusion

(6-1) Even though we are not certain the timing, it is a well formed opinion among the authorities that influenza pandemic is an incident that occurs periodically in every several decades. Once it happens, its impact to our human lives might be huge and severe. Pandemic risk is a kind of catastrophic risk. The research for the quantitative evaluation of pandemic risk has just been started.

There are several important differences between the natures of pandemic risk compared to the earthquake risk. The author understands that the most sophisticated risk evaluation model should reflect the nature of the pandemic risk. As a starting point, however, the author found that some simple assumptions about the catastrophic event, which can be applied both for pandemic risk as well as earthquake, is useful to apply it for our traditional, standard life contingency model to extend.

In this paper, the author extended the standard life contingency model with an additional mortality assumption caused by pandemic or other catastrophic event and showed some closed expression of actuarial notations such as single premium, net premium rate for continuous premium pay whole life. Further, the author considered an insurance portfolio and its mortality risk and showed a theoretical approach to calculate the standard deviation of the loss stochastic variable (L) of that portfolio, without using stochastic simulation. The standard deviation can be used as a risk measure like value at risk (VaR) or tail value at risk (TVaR). The key idea in this paper's approach is to express the joint cumulative probability distribution function of the lifetime stochastic variables of two insured in a closed form and use to derive the discrete version of joint probability density function.

In order to do some more detailed analysis regarding the distribution of the loss stochastic variable, stochastic simulation is a strong approach. One advantage of our approach is its convenience. Our approach can compliment the stochastic simulation approach.

In this paper, the author developed a prototype simulation model which does not include the nature of the pandemic risk. In order to incorporate such characteristics, the author think we need to have deeper understanding about the mathematics behind infectious diseases and epidemic modeling. For example, basic reproductive ratio, SIR (Susceptible, Infective and Removed) model and agency based modeling are some key ideas, which might be an interesting area for an actuary.

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<Appendix A> Theoretical Calculation Result of Single Premium and Some Other Actuarial Figures

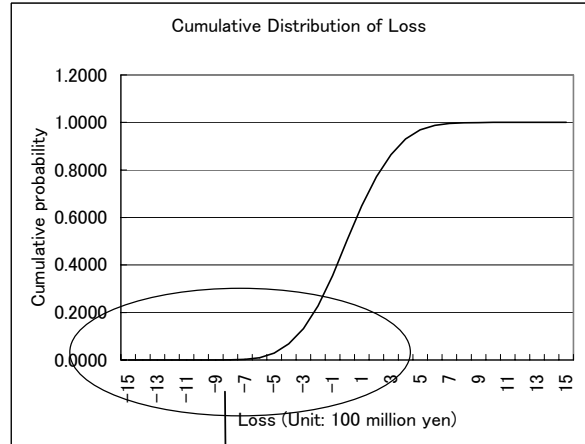
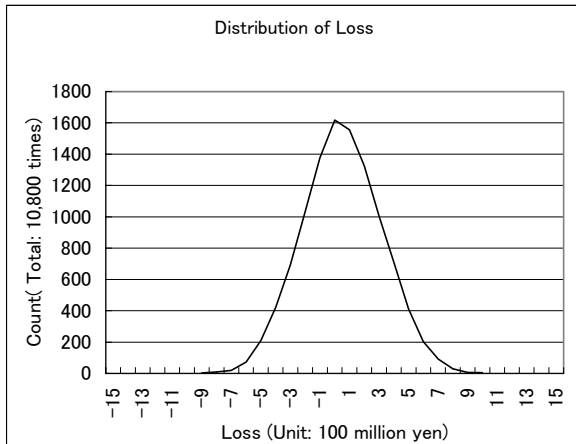
Mortality Table: 2007 Japan Mortality Table (Male) <Note> Uniform distribution of deaths in each year of age is assumed when converting discrete model into continuing model
 Assumed Interest $i = 2.00\%$
 Parameters $\lambda = 0.02$ $r = 0.98$

| Age x | Mortality q_x | Standard Model (Discrete) | | | Standard Model (Continuous) | | | Model A | | | Model B | | |
|----------|--------------------|---------------------------|---------------------|---------|-----------------------------|----------|---------|--------------|--------------|--------------|--------------|--------------|--------------|
| | | A_x | $a_{\text{due } x}$ | P_x | A_x^- | a_x^- | P_x^- | $A_x^{-(A)}$ | $a_x^{-(A)}$ | $P_x^{-(A)}$ | $A_x^{-(B)}$ | $a_x^{-(B)}$ | $P_x^{-(B)}$ |
| 0 | 0.00108 | 0.22117 | 39.72016 | 0.00557 | 0.22338 | 39.21814 | 0.00570 | 0.23249 | 38.75799 | 0.00600 | 0.22953 | 38.90754 | 0.00590 |
| 1 | 0.00075 | 0.22476 | 39.53726 | 0.00568 | 0.22700 | 39.03523 | 0.00582 | 0.23599 | 38.58101 | 0.00612 | 0.23309 | 38.72765 | 0.00602 |
| 2 | 0.00049 | 0.22868 | 39.33751 | 0.00581 | 0.23096 | 38.83548 | 0.00595 | 0.23983 | 38.38737 | 0.00625 | 0.23698 | 38.53106 | 0.00615 |
| 3 | 0.00031 | 0.23287 | 39.12343 | 0.00595 | 0.23520 | 38.62139 | 0.00609 | 0.24394 | 38.17957 | 0.00639 | 0.24116 | 38.32026 | 0.00629 |
| 4 | 0.00021 | 0.23729 | 38.89796 | 0.00610 | 0.23966 | 38.39591 | 0.00624 | 0.24828 | 37.96050 | 0.00654 | 0.24556 | 38.09818 | 0.00645 |
| 5 | 0.00017 | 0.24188 | 38.66403 | 0.00626 | 0.24429 | 38.16198 | 0.00640 | 0.25279 | 37.73306 | 0.00670 | 0.25012 | 37.86771 | 0.00661 |
| 6 | 0.00016 | 0.24659 | 38.42385 | 0.00642 | 0.24905 | 37.92179 | 0.00657 | 0.25741 | 37.49941 | 0.00686 | 0.25481 | 37.63104 | 0.00677 |
| 7 | 0.00016 | 0.25140 | 38.17843 | 0.00658 | 0.25391 | 37.67636 | 0.00674 | 0.26214 | 37.26057 | 0.00704 | 0.25960 | 37.38919 | 0.00694 |
| 8 | 0.00016 | 0.25631 | 37.92807 | 0.00676 | 0.25887 | 37.42599 | 0.00692 | 0.26697 | 37.01682 | 0.00721 | 0.26448 | 37.14244 | 0.00712 |
| 9 | 0.00015 | 0.26132 | 37.67266 | 0.00694 | 0.26392 | 37.17057 | 0.00710 | 0.27190 | 36.76806 | 0.00739 | 0.26947 | 36.89068 | 0.00730 |
| 10 | 0.00014 | 0.26644 | 37.41172 | 0.00712 | 0.26909 | 36.90963 | 0.00729 | 0.27693 | 36.51381 | 0.00758 | 0.27456 | 36.63345 | 0.00749 |
| 11 | 0.00013 | 0.27166 | 37.14516 | 0.00731 | 0.27437 | 36.64306 | 0.00749 | 0.28208 | 36.25396 | 0.00778 | 0.27977 | 36.37063 | 0.00769 |
| 12 | 0.00014 | 0.27700 | 36.87285 | 0.00751 | 0.27976 | 36.37074 | 0.00769 | 0.28734 | 35.98841 | 0.00798 | 0.28508 | 36.10211 | 0.00790 |
| 13 | 0.00018 | 0.28244 | 36.59544 | 0.00772 | 0.28526 | 36.09331 | 0.00790 | 0.29269 | 35.71776 | 0.00819 | 0.29050 | 35.82852 | 0.00811 |
| 14 | 0.00025 | 0.28796 | 36.31388 | 0.00793 | 0.29083 | 35.81175 | 0.00812 | 0.29914 | 35.44299 | 0.00841 | 0.29600 | 35.55083 | 0.00833 |
| 15 | 0.00036 | 0.29355 | 36.02917 | 0.00815 | 0.29647 | 35.52703 | 0.00834 | 0.30364 | 35.16506 | 0.00863 | 0.30156 | 35.27000 | 0.00855 |
| 16 | 0.00049 | 0.29916 | 35.74262 | 0.00837 | 0.30215 | 35.24047 | 0.00857 | 0.30918 | 34.88529 | 0.00886 | 0.30716 | 34.98735 | 0.00878 |
| 17 | 0.00062 | 0.30481 | 35.45484 | 0.00860 | 0.30785 | 34.95268 | 0.00881 | 0.31474 | 34.60426 | 0.00910 | 0.31278 | 34.70349 | 0.00901 |
| 18 | 0.00073 | 0.31048 | 35.16574 | 0.00883 | 0.31357 | 34.66357 | 0.00905 | 0.32034 | 34.32189 | 0.00933 | 0.31843 | 34.41831 | 0.00925 |
| 19 | 0.00080 | 0.31619 | 34.87451 | 0.00907 | 0.31934 | 34.37234 | 0.00929 | 0.32597 | 34.03738 | 0.00958 | 0.32412 | 34.13102 | 0.00950 |
| 20 | 0.00084 | 0.32197 | 34.57967 | 0.00931 | 0.32518 | 34.07748 | 0.00954 | 0.33168 | 33.74925 | 0.00983 | 0.32988 | 33.84013 | 0.00975 |
| 21 | 0.00086 | 0.32784 | 34.28006 | 0.00956 | 0.33111 | 33.77786 | 0.00980 | 0.33748 | 33.45636 | 0.01009 | 0.33573 | 33.54452 | 0.01001 |
| 22 | 0.00085 | 0.33383 | 33.97488 | 0.00983 | 0.33715 | 33.47267 | 0.01007 | 0.34339 | 33.15792 | 0.01036 | 0.34169 | 33.24337 | 0.01028 |
| 23 | 0.00084 | 0.33994 | 33.66299 | 0.01010 | 0.34333 | 33.16077 | 0.01035 | 0.34943 | 32.85279 | 0.01064 | 0.34779 | 32.93556 | 0.01056 |
| 24 | 0.00083 | 0.34619 | 33.34426 | 0.01038 | 0.34964 | 32.84203 | 0.01065 | 0.35561 | 32.54084 | 0.01093 | 0.35402 | 32.62095 | 0.01085 |
| 25 | 0.00082 | 0.35258 | 33.01855 | 0.01068 | 0.35609 | 32.51631 | 0.01095 | 0.36192 | 32.22193 | 0.01123 | 0.36039 | 32.29941 | 0.01116 |
| 26 | 0.00081 | 0.35910 | 32.68572 | 0.01099 | 0.36268 | 32.18347 | 0.01127 | 0.36838 | 31.89592 | 0.01155 | 0.36689 | 31.97078 | 0.01148 |
| 27 | 0.00080 | 0.36577 | 32.34563 | 0.01131 | 0.36942 | 31.84337 | 0.01160 | 0.37498 | 31.56266 | 0.01188 | 0.37355 | 31.63494 | 0.01181 |
| 28 | 0.00081 | 0.37259 | 31.99815 | 0.01164 | 0.37630 | 31.49587 | 0.01195 | 0.38172 | 31.22200 | 0.01223 | 0.38034 | 31.29173 | 0.01215 |
| 29 | 0.00083 | 0.37953 | 31.64374 | 0.01199 | 0.38332 | 31.14146 | 0.01231 | 0.38861 | 30.87443 | 0.01259 | 0.38727 | 30.94164 | 0.01252 |
| 30 | 0.00086 | 0.38662 | 31.28258 | 0.01236 | 0.39047 | 30.78028 | 0.01269 | 0.39562 | 30.52010 | 0.01296 | 0.39434 | 30.58482 | 0.01289 |
| 31 | 0.00089 | 0.39383 | 30.91482 | 0.01274 | 0.39775 | 30.41251 | 0.01308 | 0.40277 | 30.15915 | 0.01335 | 0.40154 | 30.22143 | 0.01329 |
| 32 | 0.00092 | 0.40117 | 30.54029 | 0.01314 | 0.40517 | 30.03798 | 0.01349 | 0.41005 | 29.79144 | 0.01376 | 0.40887 | 29.85129 | 0.01370 |
| 33 | 0.00096 | 0.40865 | 30.15885 | 0.01355 | 0.41272 | 29.65652 | 0.01392 | 0.41747 | 29.41678 | 0.01419 | 0.41633 | 29.47426 | 0.01413 |
| 34 | 0.00100 | 0.41626 | 29.77060 | 0.01398 | 0.42041 | 29.26826 | 0.01436 | 0.42502 | 29.03531 | 0.01464 | 0.42393 | 29.09045 | 0.01457 |
| 35 | 0.00105 | 0.42401 | 29.37539 | 0.01443 | 0.42824 | 28.87303 | 0.01483 | 0.43272 | 28.64684 | 0.01511 | 0.43167 | 28.69969 | 0.01504 |
| 36 | 0.00112 | 0.43190 | 28.97332 | 0.01491 | 0.43620 | 28.47095 | 0.01532 | 0.44055 | 28.25149 | 0.01559 | 0.43954 | 28.30209 | 0.01553 |
| 37 | 0.00119 | 0.43991 | 28.56478 | 0.01540 | 0.44429 | 28.06240 | 0.01583 | 0.44850 | 27.84964 | 0.01610 | 0.44755 | 27.89802 | 0.01604 |
| 38 | 0.00128 | 0.44805 | 28.14957 | 0.01592 | 0.45251 | 27.64718 | 0.01637 | 0.45659 | 27.44108 | 0.01664 | 0.45568 | 27.48730 | 0.01658 |
| 39 | 0.00137 | 0.45631 | 27.72806 | 0.01646 | 0.46086 | 27.22565 | 0.01693 | 0.46481 | 27.02617 | 0.01720 | 0.46394 | 27.07026 | 0.01714 |
| 40 | 0.00148 | 0.46471 | 27.30002 | 0.01702 | 0.46934 | 26.79759 | 0.01751 | 0.47316 | 26.60469 | 0.01778 | 0.47233 | 26.64671 | 0.01773 |
| 41 | 0.00161 | 0.47322 | 26.86578 | 0.01761 | 0.47794 | 26.36334 | 0.01813 | 0.48163 | 26.17695 | 0.01840 | 0.48084 | 26.21696 | 0.01834 |
| 42 | 0.00176 | 0.48185 | 26.42564 | 0.01823 | 0.48665 | 25.92319 | 0.01877 | 0.49022 | 25.74326 | 0.01904 | 0.48946 | 25.78129 | 0.01899 |
| 43 | 0.00192 | 0.49059 | 25.97988 | 0.01888 | 0.49548 | 25.47741 | 0.01945 | 0.49892 | 25.30388 | 0.01972 | 0.49820 | 25.33999 | 0.01966 |
| 44 | 0.00211 | 0.49944 | 25.52849 | 0.01956 | 0.50442 | 25.02601 | 0.02016 | 0.50773 | 24.85881 | 0.02042 | 0.50705 | 24.89305 | 0.02037 |
| 45 | 0.00231 | 0.50839 | 25.07196 | 0.02028 | 0.51346 | 24.56947 | 0.02090 | 0.51665 | 24.40853 | 0.02117 | 0.51601 | 24.44094 | 0.02111 |
| 46 | 0.00254 | 0.51745 | 24.61025 | 0.02103 | 0.52260 | 24.10774 | 0.02168 | 0.52567 | 23.95299 | 0.02195 | 0.52506 | 23.98363 | 0.02189 |
| 47 | 0.00277 | 0.52659 | 24.14378 | 0.02181 | 0.53184 | 23.64125 | 0.02250 | 0.53478 | 23.49260 | 0.02276 | 0.53421 | 23.52154 | 0.02271 |
| 48 | 0.00304 | 0.53584 | 23.67223 | 0.02264 | 0.54118 | 23.16969 | 0.02336 | 0.54400 | 23.02705 | 0.02362 | 0.54346 | 23.05433 | 0.02357 |
| 49 | 0.00333 | 0.54517 | 23.19619 | 0.02350 | 0.55061 | 22.69363 | 0.02426 | 0.55331 | 22.55692 | 0.02453 | 0.55281 | 22.58259 | 0.02448 |
| 50 | 0.00365 | 0.55459 | 22.71576 | 0.02441 | 0.56012 | 22.21318 | 0.02522 | 0.56271 | 22.08230 | 0.02548 | 0.56223 | 22.10643 | 0.02543 |
| 51 | 0.00401 | 0.56409 | 22.23122 | 0.02537 | 0.56972 | 21.72863 | 0.02622 | 0.57219 | 21.60348 | 0.02649 | 0.57175 | 21.62611 | 0.02644 |
| 52 | 0.00440 | 0.57367 | 21.74303 | 0.02638 | 0.57938 | 21.24042 | 0.02728 | 0.58175 | 21.12091 | 0.02754 | 0.58133 | 21.14211 | 0.02750 |
| 53 | 0.00480 | 0.58331 | 21.25140 | 0.02745 | 0.58912 | 20.74878 | 0.02839 | 0.59138 | 20.63479 | 0.02866 | 0.59098 | 20.65460 | 0.02861 |
| 54 | 0.00522 | 0.59302 | 20.75605 | 0.02857 | 0.59893 | 20.25342 | 0.02957 | 0.60108 | 20.14484 | 0.02984 | 0.60071 | 20.16333 | 0.02979 |
| 55 | 0.00567 | 0.60281 | 20.25692 | 0.02976 | 0.60881 | 19.75426 | 0.03082 | 0.61086 | 19.65100 | 0.03109 | 0.61052 | 19.66821 | 0.03104 |
| 56 | 0.00615 | 0.61267 | 19.75406 | 0.03101 | 0.61877 | 19.25139 | 0.03214 | 0.62071 | 19.15332 | 0.03241 | 0.62040 | 19.16932 | 0.03236 |
| 57 | 0.00666 | 0.62260 | 19.24751 | 0.03235 | 0.62880 | 18.74483 | 0.03355 | 0.63064 | 18.65184 | 0.03381 | 0.63035 | 18.66667 | 0.03377 |
| 58 | 0.00718 | 0.63260 | 18.73725 | 0.03376 | 0.63891 | 18.23455 | 0.03504 | 0.64065 | 18.14653 | 0.03530 | 0.64038 | 18.16025 | 0.03526 |
| 59 | 0.00774 | 0.64269 | 18.22284 | 0.03527 | 0.64910 | 17.72012 | 0.03663 | 0.65074 | 17.63694 | 0.03690 | 0.65049 | 17.64960 | 0.03686 |

| Age | Mortality | Standard Model (Discrete) | | | Standard Model (Continuous) | | | Model A | | | Model B | | |
|-----|-----------|---------------------------|-----------|---------|-----------------------------|----------|---------|--------------|--------------|--------------|--------------|--------------|--------------|
| x | q_x | A_x | a due x | P_x | A_x^- | a_x^- | P_x^- | $A_x^{-(A)}$ | $a_x^{-(A)}$ | $P_x^{-(A)}$ | $A_x^{-(B)}$ | $a_x^{-(B)}$ | $P_x^{-(B)}$ |
| 60 | 0.00834 | 0.65286 | 17.70433 | 0.03688 | 0.65936 | 17.20159 | 0.03833 | 0.66092 | 17.12314 | 0.03860 | 0.66069 | 17.13480 | 0.03856 |
| 61 | 0.00902 | 0.66310 | 17.18171 | 0.03859 | 0.66971 | 16.67895 | 0.04015 | 0.67118 | 16.60510 | 0.04042 | 0.67096 | 16.61581 | 0.04038 |
| 62 | 0.00981 | 0.67342 | 16.65558 | 0.04043 | 0.68013 | 16.15281 | 0.04211 | 0.68151 | 16.08342 | 0.04237 | 0.68131 | 16.09323 | 0.04234 |
| 63 | 0.01072 | 0.68379 | 16.12689 | 0.04240 | 0.69060 | 15.62410 | 0.04420 | 0.69189 | 15.55906 | 0.04447 | 0.69171 | 15.56801 | 0.04443 |
| 64 | 0.01180 | 0.69418 | 15.59663 | 0.04451 | 0.70110 | 15.09382 | 0.04645 | 0.70231 | 15.03297 | 0.04672 | 0.70215 | 15.04112 | 0.04668 |
| 65 | 0.01306 | 0.70458 | 15.06634 | 0.04677 | 0.71160 | 14.56352 | 0.04886 | 0.71273 | 14.50671 | 0.04913 | 0.71258 | 14.51411 | 0.04910 |
| 66 | 0.01452 | 0.71495 | 14.53753 | 0.04918 | 0.72208 | 14.03469 | 0.05145 | 0.72312 | 13.98177 | 0.05172 | 0.72299 | 13.98847 | 0.05168 |
| 67 | 0.01616 | 0.72526 | 14.01173 | 0.05176 | 0.73249 | 13.50887 | 0.05422 | 0.73346 | 13.45969 | 0.05449 | 0.73334 | 13.46573 | 0.05446 |
| 68 | 0.01794 | 0.73549 | 13.48996 | 0.05452 | 0.74282 | 12.98709 | 0.05720 | 0.74372 | 12.94147 | 0.05747 | 0.74362 | 12.94691 | 0.05744 |
| 69 | 0.01986 | 0.74564 | 12.97249 | 0.05748 | 0.75307 | 12.46960 | 0.06039 | 0.75390 | 12.42739 | 0.06066 | 0.75381 | 12.43226 | 0.06063 |
| 70 | 0.02193 | 0.75570 | 12.45938 | 0.06065 | 0.76323 | 11.95647 | 0.06383 | 0.76400 | 11.91752 | 0.06411 | 0.76392 | 11.92187 | 0.06408 |
| 71 | 0.02415 | 0.76567 | 11.95065 | 0.06407 | 0.77331 | 11.44772 | 0.06755 | 0.77402 | 11.41187 | 0.06783 | 0.77394 | 11.41574 | 0.06780 |
| 72 | 0.02657 | 0.77557 | 11.44608 | 0.06776 | 0.78330 | 10.94314 | 0.07158 | 0.78395 | 10.91023 | 0.07185 | 0.78388 | 10.91367 | 0.07183 |
| 73 | 0.02923 | 0.78538 | 10.94584 | 0.07175 | 0.79320 | 10.44288 | 0.07596 | 0.79380 | 10.41276 | 0.07623 | 0.79374 | 10.41579 | 0.07621 |
| 74 | 0.03223 | 0.79509 | 10.45021 | 0.07608 | 0.80302 | 9.94724 | 0.08073 | 0.80356 | 9.91976 | 0.08101 | 0.80351 | 9.92243 | 0.08098 |
| 75 | 0.03568 | 0.80470 | 9.96024 | 0.08079 | 0.81272 | 9.45724 | 0.08594 | 0.81322 | 9.43226 | 0.08622 | 0.81317 | 9.43460 | 0.08619 |
| 76 | 0.03961 | 0.81416 | 9.47760 | 0.08590 | 0.82228 | 8.97459 | 0.09162 | 0.82273 | 8.95195 | 0.09190 | 0.82269 | 8.95399 | 0.09188 |
| 77 | 0.04400 | 0.82346 | 9.00379 | 0.09146 | 0.83166 | 8.50077 | 0.09783 | 0.83207 | 8.48032 | 0.09812 | 0.83203 | 8.48208 | 0.09809 |
| 78 | 0.04877 | 0.83256 | 8.53961 | 0.09749 | 0.84085 | 8.03657 | 0.10463 | 0.84122 | 8.01816 | 0.10491 | 0.84119 | 8.01969 | 0.10489 |
| 79 | 0.05425 | 0.84148 | 8.08469 | 0.10408 | 0.84986 | 7.58164 | 0.11209 | 0.85019 | 7.56513 | 0.11238 | 0.85016 | 7.56644 | 0.11236 |
| 80 | 0.06039 | 0.85018 | 7.64091 | 0.11127 | 0.85865 | 7.13784 | 0.12030 | 0.85894 | 7.12309 | 0.12059 | 0.85892 | 7.12421 | 0.12056 |
| 81 | 0.06728 | 0.85865 | 7.20908 | 0.11911 | 0.86720 | 6.70600 | 0.12932 | 0.86746 | 6.69287 | 0.12961 | 0.86744 | 6.69382 | 0.12959 |
| 82 | 0.07500 | 0.86686 | 6.79010 | 0.12767 | 0.87550 | 6.28701 | 0.13926 | 0.87573 | 6.27536 | 0.13955 | 0.87572 | 6.27616 | 0.13953 |
| 83 | 0.08364 | 0.87481 | 6.38476 | 0.13702 | 0.88353 | 5.88165 | 0.15022 | 0.88373 | 5.87137 | 0.15052 | 0.88372 | 5.87204 | 0.15050 |
| 84 | 0.09329 | 0.88247 | 5.99378 | 0.14723 | 0.89127 | 5.49066 | 0.16232 | 0.89145 | 5.48160 | 0.16263 | 0.89144 | 5.48217 | 0.16261 |
| 85 | 0.10407 | 0.88985 | 5.61773 | 0.15840 | 0.89872 | 5.11460 | 0.17572 | 0.89887 | 5.10667 | 0.17602 | 0.89887 | 5.10713 | 0.17600 |
| 86 | 0.11609 | 0.89692 | 5.25720 | 0.17061 | 0.90586 | 4.75406 | 0.19054 | 0.90599 | 4.74713 | 0.19085 | 0.90599 | 4.74752 | 0.19083 |
| 87 | 0.12946 | 0.90367 | 4.91266 | 0.18395 | 0.91268 | 4.40950 | 0.20698 | 0.91280 | 4.40348 | 0.20729 | 0.91279 | 4.40380 | 0.20727 |
| 88 | 0.14432 | 0.91011 | 4.58441 | 0.19852 | 0.91918 | 4.08124 | 0.22522 | 0.91928 | 4.07603 | 0.22553 | 0.91928 | 4.07629 | 0.22552 |
| 89 | 0.16079 | 0.91622 | 4.27274 | 0.21443 | 0.92535 | 3.76956 | 0.24548 | 0.92544 | 3.76507 | 0.24580 | 0.92544 | 3.76528 | 0.24578 |
| 90 | 0.17900 | 0.92200 | 3.97778 | 0.23179 | 0.93119 | 3.47459 | 0.26800 | 0.93127 | 3.47073 | 0.26832 | 0.93127 | 3.47091 | 0.26831 |
| 91 | 0.19910 | 0.92746 | 3.69955 | 0.25070 | 0.93670 | 3.19636 | 0.29305 | 0.93677 | 3.19306 | 0.29338 | 0.93677 | 3.19320 | 0.29336 |
| 92 | 0.22119 | 0.93259 | 3.43806 | 0.27125 | 0.94188 | 2.93486 | 0.32093 | 0.94194 | 2.93205 | 0.32126 | 0.94194 | 2.93216 | 0.32124 |
| 93 | 0.24540 | 0.93739 | 3.19311 | 0.29357 | 0.94673 | 2.68989 | 0.35196 | 0.94678 | 2.68751 | 0.35229 | 0.94678 | 2.68760 | 0.35228 |
| 94 | 0.27184 | 0.94187 | 2.96444 | 0.31772 | 0.95126 | 2.46122 | 0.38650 | 0.95130 | 2.45921 | 0.38683 | 0.95130 | 2.45928 | 0.38682 |
| 95 | 0.30058 | 0.94604 | 2.75177 | 0.34379 | 0.95547 | 2.24855 | 0.42493 | 0.95551 | 2.24685 | 0.42527 | 0.95551 | 2.24690 | 0.42525 |
| 96 | 0.33166 | 0.94991 | 2.55470 | 0.37183 | 0.95938 | 2.05147 | 0.46765 | 0.95940 | 2.05004 | 0.46799 | 0.95940 | 2.05008 | 0.46798 |
| 97 | 0.36510 | 0.95348 | 2.37273 | 0.40185 | 0.96298 | 1.86949 | 0.51510 | 0.96300 | 1.86830 | 0.51544 | 0.96300 | 1.86834 | 0.51543 |
| 98 | 0.40085 | 0.95676 | 2.20537 | 0.43383 | 0.96629 | 1.70212 | 0.56770 | 0.96631 | 1.70113 | 0.56804 | 0.96631 | 1.70115 | 0.56803 |
| 99 | 0.43880 | 0.95976 | 2.05203 | 0.46771 | 0.96933 | 1.54878 | 0.62587 | 0.96935 | 1.54796 | 0.62621 | 0.96935 | 1.54798 | 0.62620 |
| 100 | 0.47877 | 0.96251 | 1.91211 | 0.50338 | 0.97210 | 1.40885 | 0.68999 | 0.97211 | 1.40817 | 0.69034 | 0.97211 | 1.40818 | 0.69033 |
| 101 | 0.52048 | 0.96500 | 1.78491 | 0.54064 | 0.97462 | 1.28165 | 0.76044 | 0.97463 | 1.28108 | 0.76079 | 0.97463 | 1.28110 | 0.76078 |
| 102 | 0.56359 | 0.96726 | 1.66961 | 0.57934 | 0.97690 | 1.16635 | 0.83758 | 0.97691 | 1.16588 | 0.83792 | 0.97691 | 1.16588 | 0.83792 |
| 103 | 0.60761 | 0.96931 | 1.56504 | 0.61935 | 0.97897 | 1.06178 | 0.92201 | 0.97898 | 1.06139 | 0.92236 | 0.97898 | 1.06140 | 0.92235 |
| 104 | 0.65200 | 0.97120 | 1.46880 | 0.66122 | 0.98088 | 0.96554 | 1.01589 | 0.98089 | 0.96522 | 1.01623 | 0.98089 | 0.96523 | 1.01622 |
| 105 | 0.69612 | 0.97306 | 1.37408 | 0.70815 | 0.98276 | 0.87081 | 1.12855 | 0.98276 | 0.87056 | 1.12888 | 0.98276 | 0.87057 | 1.12887 |
| 106 | 0.73925 | 0.97538 | 1.25564 | 0.77680 | 0.98510 | 0.75236 | 1.30935 | 0.98510 | 0.75219 | 1.30964 | 0.98510 | 0.75220 | 1.30964 |
| 107 | 1.00000 | 0.98039 | 1.00000 | 0.98039 | 0.99016 | 0.49672 | 1.99342 | 0.99017 | 0.49665 | 1.99369 | 0.99017 | 0.49665 | 1.99369 |

<Appendix B> Simulation Result

(SIM1) Standard Model, Simulation Result



Simulation Assumptions and Result

N: 21,000

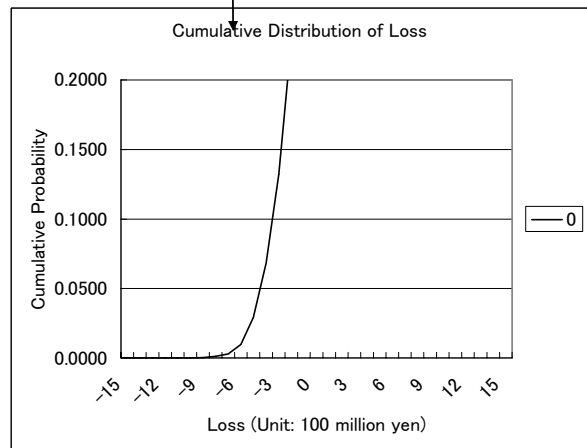
Insurance Amount: 10 million yen

Simulation: 10,800 times

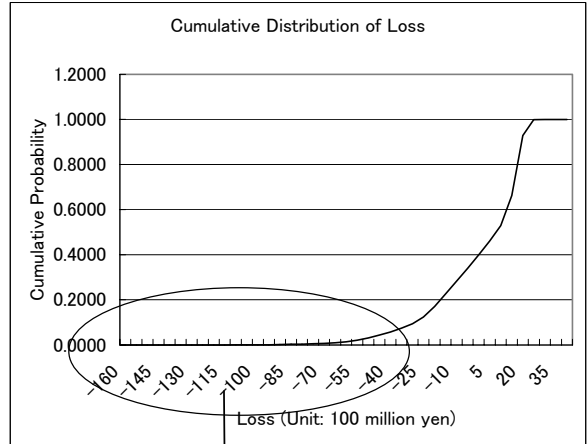
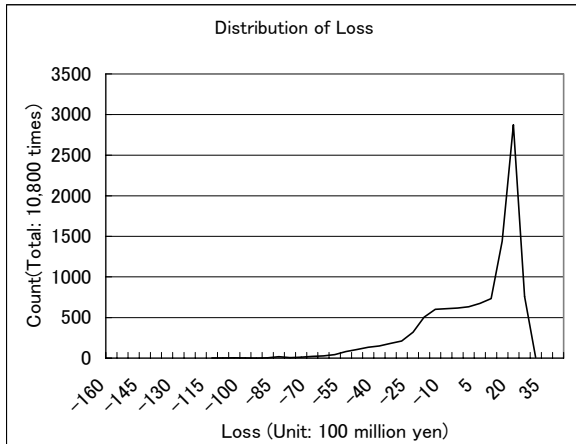
| (Unit: yen) | |
|--------------------|--------------|
| Maximum of L | 997,205,623 |
| Minimum of L | -945,271,795 |
| Average of L | -0 |
| Median of L | -3,240,661 |
| Std Deviation of L | 266,652,905 |
| Left 0.1% point | -818,872,070 |
| Left 1% point | -597,844,627 |
| Left 5% point | -434,962,284 |
| Left 10% point | -340,204,438 |

Per policy, per insurance amount

| | |
|--------------------|-------------|
| Maximum of L | 0.00474860 |
| Minimum of L | -0.00450129 |
| Average of L | -0.00000000 |
| Median of L | -0.00001543 |
| Std Deviation of L | 0.00126978 |
| Left 0.1% point | -0.00389939 |
| Left 1% point | -0.00284688 |
| Left 5% point | -0.00207125 |
| Left 10% point | -0.00162002 |



(SIM2) Model A, Simulation Result



Simulation Assumptions and Result

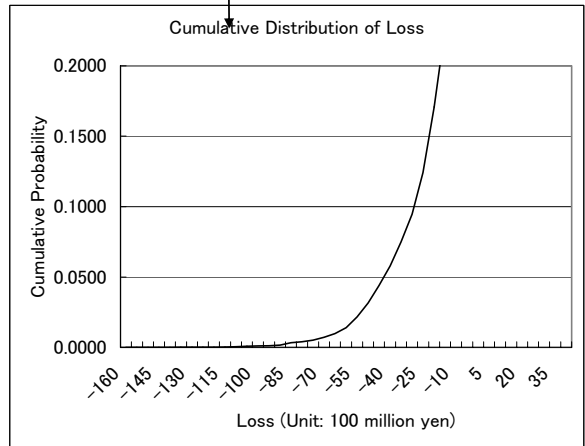
N: 21,000
 Insurance Amount: 10 million yen
 Simulation: 10,800 times

(Unit: yen)

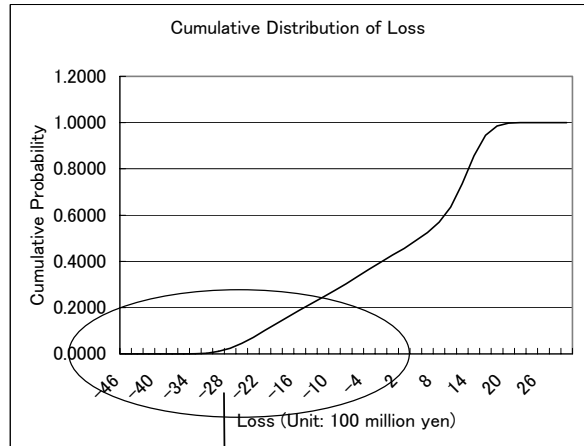
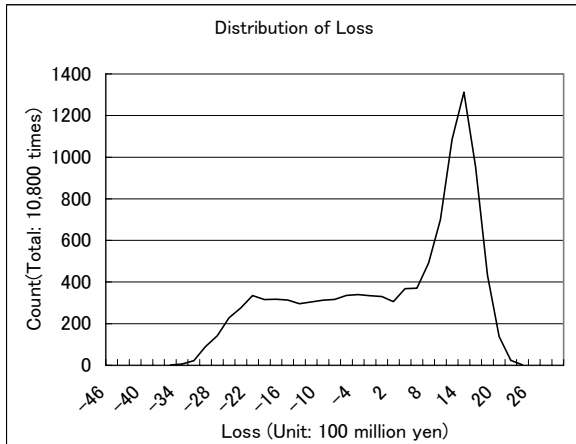
| | |
|--------------------|-----------------|
| Maximum of L | 2,742,172,806 |
| Minimum of L | -13,812,581,352 |
| Average of L | -0 |
| Median of L | 791,045,591 |
| Std Deviation of L | 2,118,330,509 |
| Left 0.1% point | -10,159,553,993 |
| Left 1% point | -6,492,448,616 |
| Left 5% point | -4,258,699,583 |
| Left 10% point | -2,887,088,157 |

Per policy, per insurance amount

| | |
|--------------------|-------------|
| Maximum of L | 0.01305797 |
| Minimum of L | -0.06577420 |
| Average of L | -0.00000000 |
| Median of L | 0.00376688 |
| Std Deviation of L | 0.01008729 |
| Left 0.1% point | -0.04837883 |
| Left 1% point | -0.03091642 |
| Left 5% point | -0.02027952 |
| Left 10% point | -0.01374804 |



(SIM3) Model B1, Simulation Result



Simulation Assumptions and Result

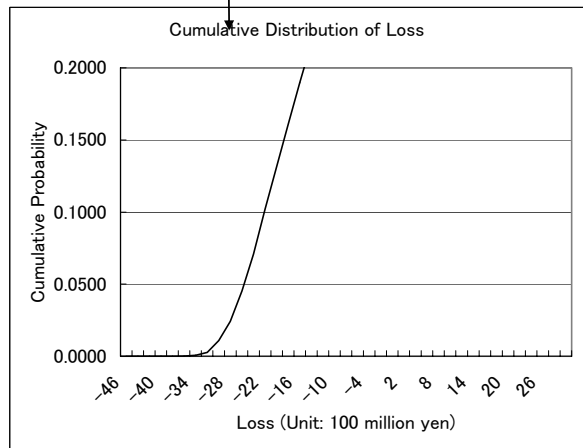
N: 21,000
 Insurance Amount: 10 million yen
 Simulation: 10,800 times

(Unit: yen)

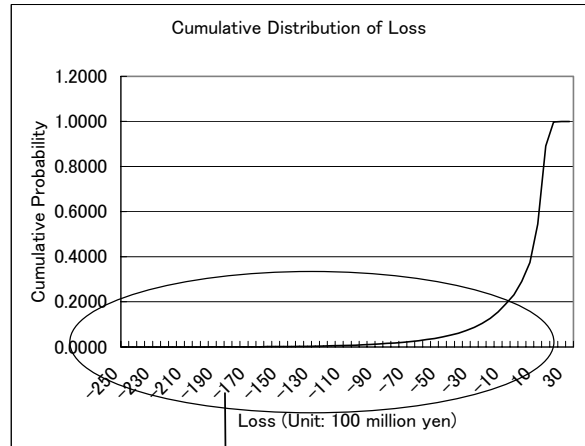
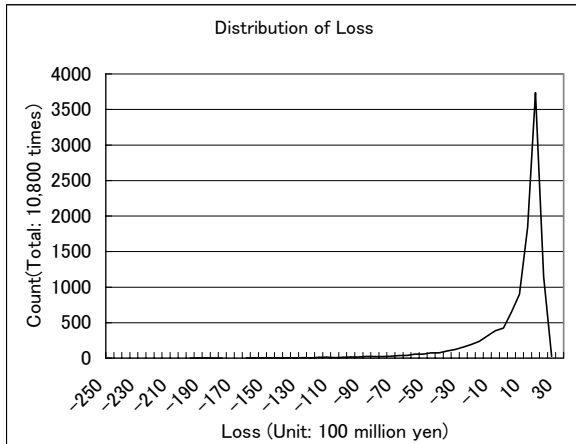
| | |
|--------------------|----------------|
| Maximum of L | 2,293,351,849 |
| Minimum of L | -3,670,193,204 |
| Average of L | -0 |
| Median of L | 459,454,609 |
| Std Deviation of L | 1,411,713,605 |
| Left 0.1% point | -3,320,663,071 |
| Left 1% point | -3,024,352,806 |
| Left 5% point | -2,558,994,237 |
| Left 10% point | -2,210,155,318 |

Per policy, per insurance amount

| | |
|--------------------|-------------|
| Maximum of L | 0.01092072 |
| Minimum of L | -0.01747711 |
| Average of L | -0.00000000 |
| Median of L | 0.00218788 |
| Std Deviation of L | 0.00672245 |
| Left 0.1% point | -0.01581268 |
| Left 1% point | -0.01440168 |
| Left 5% point | -0.01218569 |
| Left 10% point | -0.01052455 |



(SIM4) Model B, Simulation Result



Simulation Assumptions and Result

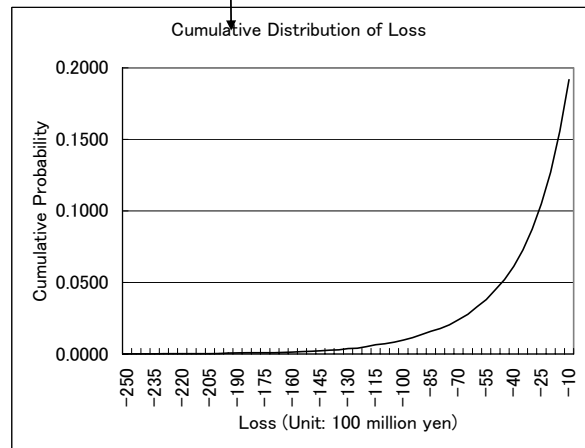
N:21,000
 Insurance Amount: 10 million yen
 Simulation: 10,800 times

(Unit: yen)

| | |
|--------------------|-----------------|
| Maximum of L | 2,285,371,246 |
| Minimum of L | -27,342,931,211 |
| Average of L | -0 |
| Median of L | 906,498,429 |
| Std Deviation of L | 2,358,719,798 |
| Left 0.1% point | -18,100,816,082 |
| Left 1% point | -9,950,098,865 |
| Left 5% point | -4,647,559,335 |
| Left 10% point | -2,618,584,083 |

Per policy, per insurance amount

| | |
|--------------------|-------------|
| Maximum of L | 0.01088272 |
| Minimum of L | -0.13020443 |
| Average of L | -0.00000000 |
| Median of L | 0.00431666 |
| Std Deviation of L | 0.01123200 |
| Left 0.1% point | -0.08619436 |
| Left 1% point | -0.04738142 |
| Left 5% point | -0.02213123 |
| Left 10% point | -0.01246945 |



<Appendix C> Standard Deviation of Loss of an Insurance Portfolio

Here we assume that n is large enough so that we can use the following approximation.

$$Std\left[\frac{Port L^*}{n}\right] \approx \left(1 + \frac{\bar{P}_x^*}{\delta}\right) \sqrt{Cov[Z^*_1, Z^*_2]}$$

Right side value has calculated using the discrete version of probability density function in the paper

| Age | Model A | Model B1 | Model B2 |
|-----|---------|----------|----------|
| 0 | 0.01167 | 0.00655 | 0.01206 |
| 1 | 0.01163 | 0.00656 | 0.01204 |
| 2 | 0.01159 | 0.00656 | 0.01202 |
| 3 | 0.01154 | 0.00657 | 0.01200 |
| 4 | 0.01150 | 0.00657 | 0.01198 |
| 5 | 0.01145 | 0.00658 | 0.01195 |
| 6 | 0.01141 | 0.00658 | 0.01193 |
| 7 | 0.01136 | 0.00659 | 0.01191 |
| 8 | 0.01131 | 0.00659 | 0.01188 |
| 9 | 0.01126 | 0.00660 | 0.01186 |
| 10 | 0.01121 | 0.00660 | 0.01183 |
| 11 | 0.01116 | 0.00660 | 0.01181 |
| 12 | 0.01111 | 0.00661 | 0.01178 |
| 13 | 0.01106 | 0.00661 | 0.01175 |
| 14 | 0.01101 | 0.00661 | 0.01173 |
| 15 | 0.01096 | 0.00662 | 0.01170 |
| 16 | 0.01090 | 0.00662 | 0.01167 |
| 17 | 0.01085 | 0.00662 | 0.01164 |
| 18 | 0.01079 | 0.00662 | 0.01161 |
| 19 | 0.01074 | 0.00662 | 0.01157 |
| 20 | 0.01068 | 0.00662 | 0.01154 |
| 21 | 0.01062 | 0.00662 | 0.01151 |
| 22 | 0.01056 | 0.00662 | 0.01147 |
| 23 | 0.01050 | 0.00662 | 0.01144 |
| 24 | 0.01044 | 0.00662 | 0.01140 |
| 25 | 0.01037 | 0.00661 | 0.01136 |
| 26 | 0.01031 | 0.00661 | 0.01132 |
| 27 | 0.01024 | 0.00661 | 0.01128 |
| 28 | 0.01018 | 0.00660 | 0.01124 |
| 29 | 0.01011 | 0.00659 | 0.01120 |
| 30 | 0.01004 | 0.00659 | 0.01115 |
| 31 | 0.00997 | 0.00658 | 0.01111 |
| 32 | 0.00990 | 0.00657 | 0.01106 |
| 33 | 0.00983 | 0.00656 | 0.01101 |
| 34 | 0.00975 | 0.00655 | 0.01096 |
| 35 | 0.00968 | 0.00654 | 0.01091 |
| 36 | 0.00960 | 0.00653 | 0.01086 |
| 37 | 0.00953 | 0.00652 | 0.01080 |
| 38 | 0.00945 | 0.00650 | 0.01074 |
| 39 | 0.00937 | 0.00649 | 0.01069 |
| 40 | 0.00929 | 0.00647 | 0.01063 |
| 41 | 0.00921 | 0.00646 | 0.01057 |
| 42 | 0.00912 | 0.00644 | 0.01050 |
| 43 | 0.00904 | 0.00642 | 0.01044 |
| 44 | 0.00895 | 0.00640 | 0.01037 |
| 45 | 0.00886 | 0.00637 | 0.01031 |
| 46 | 0.00878 | 0.00635 | 0.01024 |
| 47 | 0.00869 | 0.00633 | 0.01017 |
| 48 | 0.00859 | 0.00630 | 0.01009 |
| 49 | 0.00850 | 0.00627 | 0.01002 |
| 50 | 0.00841 | 0.00624 | 0.00994 |
| 51 | 0.00831 | 0.00621 | 0.00986 |
| 52 | 0.00821 | 0.00618 | 0.00978 |
| 53 | 0.00811 | 0.00614 | 0.00969 |
| 54 | 0.00801 | 0.00611 | 0.00960 |
| 55 | 0.00791 | 0.00607 | 0.00952 |
| 56 | 0.00781 | 0.00603 | 0.00942 |
| 57 | 0.00770 | 0.00598 | 0.00933 |
| 58 | 0.00760 | 0.00594 | 0.00923 |
| 59 | 0.00749 | 0.00589 | 0.00913 |

| Age | Model A | Model B1 | Model B2 |
|-----|---------|----------|----------|
| 60 | 0.00738 | 0.00584 | 0.00903 |
| 61 | 0.00727 | 0.00579 | 0.00892 |
| 62 | 0.00715 | 0.00574 | 0.00881 |
| 63 | 0.00704 | 0.00568 | 0.00870 |
| 64 | 0.00692 | 0.00562 | 0.00859 |
| 65 | 0.00680 | 0.00556 | 0.00847 |
| 66 | 0.00668 | 0.00550 | 0.00835 |
| 67 | 0.00656 | 0.00543 | 0.00823 |
| 68 | 0.00644 | 0.00537 | 0.00811 |
| 69 | 0.00631 | 0.00530 | 0.00798 |
| 70 | 0.00619 | 0.00522 | 0.00785 |
| 71 | 0.00606 | 0.00515 | 0.00771 |
| 72 | 0.00593 | 0.00507 | 0.00757 |
| 73 | 0.00580 | 0.00499 | 0.00743 |
| 74 | 0.00567 | 0.00490 | 0.00729 |
| 75 | 0.00554 | 0.00482 | 0.00714 |
| 76 | 0.00540 | 0.00473 | 0.00699 |
| 77 | 0.00527 | 0.00464 | 0.00684 |
| 78 | 0.00513 | 0.00454 | 0.00668 |
| 79 | 0.00499 | 0.00444 | 0.00653 |
| 80 | 0.00485 | 0.00434 | 0.00636 |
| 81 | 0.00471 | 0.00424 | 0.00620 |
| 82 | 0.00457 | 0.00414 | 0.00604 |
| 83 | 0.00443 | 0.00403 | 0.00587 |
| 84 | 0.00429 | 0.00392 | 0.00570 |
| 85 | 0.00415 | 0.00381 | 0.00552 |
| 86 | 0.00401 | 0.00370 | 0.00535 |
| 87 | 0.00387 | 0.00359 | 0.00518 |
| 88 | 0.00372 | 0.00347 | 0.00500 |
| 89 | 0.00358 | 0.00335 | 0.00482 |
| 90 | 0.00344 | 0.00323 | 0.00464 |
| 91 | 0.00330 | 0.00311 | 0.00446 |
| 92 | 0.00316 | 0.00299 | 0.00428 |
| 93 | 0.00302 | 0.00287 | 0.00410 |
| 94 | 0.00288 | 0.00275 | 0.00392 |
| 95 | 0.00274 | 0.00262 | 0.00373 |
| 96 | 0.00260 | 0.00249 | 0.00355 |
| 97 | 0.00246 | 0.00237 | 0.00336 |
| 98 | 0.00231 | 0.00224 | 0.00317 |
| 99 | 0.00217 | 0.00210 | 0.00298 |
| 100 | 0.00203 | 0.00197 | 0.00279 |
| 101 | 0.00188 | 0.00183 | 0.00259 |
| 102 | 0.00174 | 0.00169 | 0.00239 |
| 103 | 0.00159 | 0.00155 | 0.00219 |
| 104 | 0.00142 | 0.00139 | 0.00197 |
| 105 | 0.00124 | 0.00122 | 0.00172 |
| 106 | 0.00096 | 0.00095 | 0.00133 |
| 107 | 0.00000 | 0.00000 | 0.00000 |